





Overview

'Classical' sample size estimation in BE Patient's & producer's risk Power in study planning Uncertainties Variability Test/Reference-ratio Sensitivity analysis Recent developments Review of guidelines



α and β

All formal decisions are subjected to two types of error:

- α Probability of Error Type I (aka Risk Type I)
- β Probability of Error Type II (aka Risk Type II) Example from the justice system:

Verdict	Defendant innocent	Defendant guilty
Presumption of innocence not accepted (guilty)	Error type I	Correct
Presumption of innocence accepted (not guilty)	Correct	Error type II





•Or in more statistical terms:

Decision	Null hypothesis true	Null hypothesis false				
Null hypothesis rejected	Error type I	Correct (<i>H</i> _a)				
Failed to reject null hypothesis	Correct (H ₀)	Error type II				

•In BE-testing the null hypothesis is bioinequivalence $(\mu_1 \neq \mu_2)!$

Decision	Null hypothesis true	Null hypothesis false
Null hypothesis rejected	Patient's risk	Correct (BE)
Failed to reject null hypothesis	Correct (not BE)	Producer's risk



Cl

• Patient's Risk to be treated with an inequivalent formulation (H_0 falsely rejected)

- BA of the test compared to reference in a particular patient is risky <u>either</u> below 80% <u>or</u> above 125%.
- If we keep the risk of particular patients at α 0.05 (5%), the risk of the entire population of patients (<80% and >125%) is 2×α(10%) expressed as: 90% CI = 1 2×α = 0.90.





... and β

• Producer's Risk to get no approval of an equivalent formulation (H_0 falsely not rejected)

Set in study planning to ≤ 0.2 (20%), where power = $1 - \beta = \geq 80\%$

If power is set to 80 %,

one out of five studies will fail just by chance!

$$\alpha$$
 0.05BEnot BEβ 0.20

A posteriori (post hoc) power does not make sense! Either a study has demonstrated BE or not.



Power Curves

Power to show BE with 12 - 36 subjects for CV_{intra} 20%

<mark>n</mark> power	24 0.896	 ↓ 16: → 0.735
μ _T /μ _R	1.05	↓ 1.10:
power	0.903	→ 0.700





Power vs. Sample Size

- It is not possible to calculate the required sample size directly.
- Power is calculated instead; the smallest sample size which fulfills the minimum target power is used.
 - Example: α 0.05, target power 80% (β 0.2), T/R 0.95, CV_{intra} 20% \rightarrow minimum sample size 19 (power 81%), rounded *up* to the next even number in a 2×2 study (power 83%).





Power vs. Sample Size

2×2 cross-over, T/R 0.95, AR 80–125%, target power 80%

sample size — power — power for n=12





Background

 Reminder: Sample Size is not directly obtained – only power

- Solution given by DB Owen (1965) as a difference of two bivariate noncentral *t*-distributions
 - Definite integrals cannot be solved in closed form
 - 'Exact' methods rely on numerical methods (currently the most advanced is AS 243 of RV Lenth; implemented in R, FARTSSIE, EFG). nQuery uses an earlier version (AS 184).



Background

Power estimations...

- 'Brute force' methods (also called 'resampling' or 'Monte Carlo') converge asymptotically to the true power; need a good random number generator (e.g., Mersenne Twister) and may be time-consuming
- 'Asymptotic' methods use large sample approximations
- Approximations provide algorithms which should converge to the desired power based on the *t*-distribution



Sample Size (Guidelines)

Recommended minimum

- 12 WHO, EU, CAN, NZ, AUS, AR, MZ, ASEAN States, RSA, Russia (?)
- USA 'A pilot study that documents BE can be appropriate, provided its design and execution are suitable and a sufficient number of subjects (e.g., 12) have completed the study.'
- **18** Russia (2008)
- 20 RSA (MR formulations)
- **24** Saudia Arabia (12 to 24 if statistically justifiable)
- **24** Brazil
- Sufficient number' Japan



Sample Size (Limits)

•Maximum

NZ: If the calculated number of subjects appears to be higher than is ethically justifiable, it may be necessary to accept a statistical power which is less than desirable. Normally it is not practical to use more than about 40 subjects in a bioavailability study.

All others: Not specified (judged by IEC/IRB and/or local Authorities).

ICH E9, Section 3.5 applies:

The number of subjects in a clinical trial should always be large enough to provide a reliable answer to the questions addressed.



Power & Sample Size

Reminder

- Generally power is set to at least 80% (β, error type II: producers's risk to get no approval for a bioequivalent formulation; power = 1 β).
 1 out of 5 studies will fail just by chance!
- If you plan for power of less than 70%, probably you will face problems with the ethics committee (ICH E9).
- If you plan for power of more than 90% (especially with low variability drugs), problems with regulators are possible ('forced bioequivalence').
- Add subjects ('alternates') according to the expected drop-out rate especially for studies with more than two periods or multiple-dose studies.



US FDA, Canada TPD

- Statistical Approaches to Establishing Bioequivalence (2001)
 - Based on maximum difference of 5%.
 - Sample size based on 80 90% power.
- •Draft GL (2010)*
 - Consider potency differences.
 - Sample size based on 80 90% power.
 - Do not interpolate linear between CVs (as stated in the GL)!
 - * All points removed in current (2012) GL.





•EMEA NfG on BA/BE (2001)

Detailed information (data sources, significance level, expected deviation of test from reference, desired power).

•EMA GL on BE (2010)

Batches must not differ more than 5% in actual content.

The number of subjects to be included in the study should be based on an appropriate sample size calculation.

Cookbook?

Information



Hierarchy of Designs

•The more 'sophisticated' a design is, the more information can be extracted.

Hierarchy of designs:
 Fully replicate (TRTR | RTRT, TRT | RTR) →
 Partial replicate (TRR | RTR | RRT) →
 Standard 2×2 cross-over (RT | RT) →
 Parallel (R | T)

Variances which can be estimated:

Parallel: 2×2 Xover: Partial replicate: Full replicate:

total variance (between + within)

- + between, within subjects 🕩
- + within subjects (reference) 🕩
- + within subjects (reference, test) 🕩



Coefficient(s) of Variation

- From any design one gets variances of *lower* design levels as well.
 - Total CV% from a 2×2 cross-over used in planning a parallel design study:
 - Intra-subject CV% (within)
 - Inter-subject CV% (between)
 - Total CV% (pooled)

$$\sim CV_{\text{intra}} \% = 100 \cdot \sqrt{e^{MSE_W} - 1}$$

$$\sim CV_{\text{inter}} \% = 100 \cdot \sqrt{e^{\frac{MSE_B - MSE_W}{2}} - 1}$$

$$CV_{\text{total}} \% = 100 \cdot \sqrt{e^{\frac{MSE_B + MSE_W}{2}}} - 1$$



Coefficient(s) of Variation

However, CVs of higher design levels not available.

If only mean ± SD of reference is available...

- Avoid 'rule of thumb' CV_{intra} = 60% of CV_{total}
- Don't plan a cross-over based on CV_{total}
- Examples (cross-over studies)

drug, formulation	design	n	metric	CV _{intra}	CV _{inter}	CV _{total}
methylphenidate MR	SD	12	AUC _t	7.00	19.1	20.4
paroxetine MR	MD	32	AUC _τ	25.2	55.1	62.1
lansoprazole DR	SD	47	C _{max}	47.0	25.1	54.6

- Pilot study unavoidable, unless
- Two-stage sequential design is used



Data from Pilot Studies

 Estimated CVs have a high degree of uncertainty (in the pivotal study it is more likely that you will be able to reproduce the PE, than the CV)

- The smaller the size of the pilot, the more uncertain the outcome.
- The more formulations you have tested, lesser degrees of freedom will result in worse estimates.
- Remember: CV is an estimate not carved in stone!





Pilot Studies: Sample Size

Small pilot studies (sample size <12) Are useful in checking the sampling schedule and the appropriateness of the analytical method, but are not suitable for the purpose of sample size planning! Sample sizes (T/R 0.95, power ≥80%) based on a n = 10 pilot study

library(PowerTOST)
expsampleN.TOST(alpha=0.05,
 targetpower=0.80, theta1=0.80,
 theta2=1.25, theta0=0.95, CV=0.40,
 dfCV=24-2, alpha2=0.05, design="2x2")

CV/0/		CV	ratio				
	fixed	uncertain	uncert./fixed				
20	20	24	1.200				
25	28	36	1.286				
30	40	52	1.300				
35	52	68	1.308				
40	66	86	1.303				

If pilot n=24: n=72, ratio 1.091



Pilot Studies: Sample Size

- Moderate sized pilot studies (sample size ~12–24) lead to more consistent results (both CV and PE).
 - If you stated a procedure in your protocol, even BE may be claimed in the pilot study, and no further study will be necessary (US-FDA).
 - If you have some previous hints of high intra-subject variability (>30%), a pilot study size of at least 24 subjects is reasonable. If you want reference-scaling, the pilot study has to be in a replicate design.
 - A Sequential Design avoids an unnecessarily large pilot study.



Pilot Studies: Sample Size

- Do not use the pilot study's CV, but calculate an upper confidence interval!
 - Gould (1995) recommends a 75% CI (*i.e.*, a producer's risk of 25%).
 - Apply Bayesian Methods (Julious and Owen 2006, Julious 2010) implemented in *R's PowerTOST/expsampleN.TOST*.

Unless you are under time pressure, a Two-Stage Sequential Design will help in dealing with the uncertain estimate from the pilot study.



Hints

Literature search for CV%

- Preferably other BE studies (the bigger, the better!)
- PK interaction studies (Cave: Mainly in steady state! Generally lower CV than after SD).
- Food studies (CV higher/lower than fasted!)
- If CV_{intra} not given (quite often), a little algebra helps. All you need is the 90% geometric confidence interval and the sample size.



Calculation of CV_{intra} from CI

Point estimate (PE) from the Confidence Limits

 $PE = \sqrt{CL_{lo} \cdot CL_{hi}}$

Estimate the number of subjects / sequence (example 2×2 cross-over)

If total sample size (N) is an even number, assume (!)

$$n_1 = n_2 = \frac{1}{2}N$$

If N is an odd number, assume (!)

 $n_1 = \frac{1}{2}N + \frac{1}{2}$, $n_2 = \frac{1}{2}N - \frac{1}{2}$ (not $n_1 = n_2 = \frac{1}{2}N!$)

Difference between one CL and the PE in log-scale; use the CL which is given with more significant digits

$$\Delta_{CL} = \ln PE - \ln CL_{lo} \quad or \quad \Delta_{CL} = \ln CL_{hi} - \ln PE$$



Calculation of CV_{intra} from CI (cont'd)

Calculate the Mean Square Error (MSE)



CV_{intra} from MSE as usual $CV_{intra} \% = 100 \cdot \sqrt{e^{MSE} - 1}$







Proof: CI from calculated values

Example: 90% CI [0.91 – 1.15], N 21 (n₁ = 11, n₂ = 10)

 $\ln PE = \ln \sqrt{CL_{lo} \cdot CL_{hi}} = \ln \sqrt{0.91 \times 1.15} = 0.02274$

$$SE_{\Delta} = \sqrt{\frac{2 \cdot MSE}{N}} = \sqrt{\frac{2 \times 0.04798}{21}} = 0.067598$$

$$CI = e^{\ln PE \pm t \cdot SE_{\Delta}} = e^{0.02274 \pm 1.729 \times 0.067598}$$

$$CI_{lo} = e^{0.02274 - 1.729 \times 0.067598} = 0.91$$

$$CI_{hi} = e^{0.02274 + 1.729 \times 0.067598} = 1.15$$



Sensitivity to Imbalance

If the study was more imbalanced than assumed, the estimated CV is conservative

Example: 90% CI [0.89 – 1.15], N 24 (n₁ = 16, n₂ = 8, but not reported as such); CV 24.74% in the study





No Algebra...

 Implemented in *R*-package *PowerTOST*, function *CVfromCI* (not only 2×2 cross-over, but also parallel groups, higher order cross-overs, replicate designs).
 Example:

library(PowerTOST)
CVfromCI(lower=0.91, upper=1.15, n=21, design="2x2", alpha=0.05)
[1] 0.2219886



Literature data



Doxicycline (37 studies from Blume/Mutschler, *Bioäquivalenz: Qualitätsbewertung wirkstoffgleicher Fertigarzneimittel*, GOVI-Verlag, Frankfurt am Main/Eschborn, 1989-1996)



- Intra-subject CV from different studies can be pooled (LA Gould 1995, Patterson and Jones 2006)
 - In the parametric model of log-transformed data, additivity of variances (not of CVs!) apply.
 - Do not use the arithmetic mean (or the geometric mean either) of CVs.
 - Before pooling variances must be weighted acccording to the studies' sample size and sequences
 - Larger studies are more influentual than smaller ones.
 - More sequences (with the same n) give higher CV.



• Intra-subject CV from different Xover studies • Calculate the variance from CV $\sigma_W^2 = \ln(CV_{intra}^2 + 1)$ • Calculate the total variance weighted by df $\sum \sigma_W^2 df$ • Calculate the pooled CV from total variance $CV = \sqrt{e^{\sum \sigma_W^2 df} - 1}$

• Optionally calculate an upper $(1-\alpha)$ % confidence limit on the pooled CV (recommended $\alpha = 0.25$)

$$CL_{CV} = \sqrt{e^{\sum \sigma_W^2 df / \chi^2_{\alpha, \sum df}} - 1}$$



Degrees of freedom of various Xover designs

Name	df	Name in PowerTOST
2×2×2 cross over	n – 2	2x2
3×3 Latin Squares	2n – 4	3x3
6 sequence Williams' design	2n – 4	3x6x3
4×4 Latin Squares, Williams' design	3n – 6	4x4
2×2×3 fully replicated design	2n – 3	2x2x3
2×2×4 fully replicated design	3n – 4	2x2x4
2×4×4 fully replicated design	3n – 4	2x4x4
2×3×3 partial replicate design	2n – 3	2x3x3



• Example: 3 studies, different Xover designs





R package *PowerTost* function *CVpooled*, example's data.

```
library(PowerTOST)
CVS <- ("
  PKmetric | CV | n | design |
                                  source
     AUC | 0.15 | 12 | 3x6x3
                                  study 1
                               AUC | 0.25 | 16 | 2x2 |
                                  study 2
     AUC | 0.20 | 24 |
                         2x2
                                  study 3
")
txtcon <- textConnection(CVs)</pre>
CVdata <- read.table(txtcon, header=TRUE, sep="|",
            strip.white=TRUE. as.is=TRUE)
close(txtcon)
CVsAUC <- subset(CVdata,PKmetric=="AUC")</pre>
print(CVpooled(CVsAUC, alpha=0.25), digits=4, verbose=TRUE)
Pooled CV = 0.1981 with 56 degrees of freedom
Upper 75% confidence limit of CV = 0.2131
```



Or you may combine pooling with an estimated sample size based on uncertain CVs (we will see later what that means). *R* package *PowerTost* function *expsampleN.TOST*, data of last example.
CVs and degrees of freedom must be given as vectors:
CV = c(0.15,0.25,0.2), dfCV = c(20,14,22)



library(PowerTOST)
expsampleN.TOST(alpha=0.05,
 targetpower=0.8, theta0=0.95,
 CV=c(0.15,0.25,0.2),
 dfCV=c(20,14,22),
 alpha2=0.25, design="2x2",
 print=TRUE, details=TRUE)

++++++ Equivalence test - TOST +++++++ Sample size est. with uncertain CV

```
Study design: 2x2 crossover
Design characteristics:
df = n-2, design const. = 2, step = 2
log-transformed data (multiplicative model)
alpha = 0.05, target power = 0.8
BE margins = 0.8 \dots 1.25
Null (true) ratio = 0.95
Variability data
  CV df
0.15 20
0.25 14
0.20 22
CV(pooled)
            = 0.1981467 with 56 df
one-sided upper CL = 0.2131329 (level = 75%)
Sample size search
     exp. power
n
16
   0.733033
18
  0.788859
20
  0.832028
```



• 'Doing the maths' is just the first part of the job!

Does it make sense to pool studies of different 'quality'?

- The reference product may have been subjected to many (minor only?) changes from the formulation used in early publications.
- Different bioanalytical methods are applied. Newer (e.g. LC/MS-MS) methods are not necessarily better in terms of CV (matrix effects!).
- Generally we have insufficient information about the clinical setup (e.g. posture control).
- Review studies critically; don't try to mix oil with water.



Tools

- Sample Size Tables (Phillips, Diletti, Hauschke, Chow, Julious, ...)
- Approximations (Diletti, Chow, Julious, ...)
- General purpose (SAS, S+, R, StaTable, ...)
- Specialized Software (nQuery Advisor, PASS, FARTSSIE, StudySize, ...)
- •Exact method (Owen implemented in *R*-package *PowerTOST*)*
 - * Thanks to Detlew Labes!



Approximations obsolete

Exact sample size tables still useful in checking plausibility of software's results

Approximations based on noncentral *t* (FARTSSIE17)



http://individual.utoronto.ca/ddubins/FARTSSIE17.xls

or $(\mathbb{R} / \mathbb{S} + \rightarrow$

Exact method (Owen) in *R*-package *PowerTOST*

http://cran.r-project.org/web/packages/PowerTOST/

library(PowerTOST)
 sampleN.TOST(alpha=0.05,
 targetpower=0.80, theta0=0.95,
 CV=0.30, design='2x2')

```
alpha
        <- 0.05
                    # alpha
        <- 0.30
                    # intra-subject CV
CV
                    # lower acceptance limit
theta1 <-0.80
theta2 <- 1/theta1 # upper acceptance limit
                    # expected ratio T/R
theta0 <- 0.95
                    # minimum power
PwrNeed <- 0.80
                    # Upper Limit for Search
Limit
        <- 1000
        <- 4
                    # start value of sample size search
n
        <- sqrt(2)*sqrt(log(CV^2+1))
S
repeat{
        <- qt(1-alpha,n-2)
  t
        <- sqrt(n)*(log(theta0)-log(theta1))/s
 nc1
        <- sqrt(n)*(log(theta0)-log(theta2))/s
 nc2
 prob1 <- pt(+t,n-2,nc1); prob2 <- pt(-t,n-2,nc2)
  power <- prob2-prob1
                    # increment sample size
        <- n+2
  n
  if(power >= PwrNeed | (n-2) >= Limit) break }
       <- n-2
Total
if(Total == Limit){
  cat('Search stopped at Limit', Limit,
        obtained Power', power*100, '\otimes n')
  } else
  cat('Sample Size', Total, '(Power', power*100, '%)\n')
```



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Com	oarso	

			CV	%										
original values	Method	Algorithm	5	7.5	10	12	12.5	14	15	16	17.5	18	20	22
PowerTOST 1.1-02 (2013) exact	Owen's Q	4	6	8	8	10	12	12	14	16	16	20	22
Patterson & Jones (2006)	noncentr. t	AS 243	4	5	7	8	9	11	12	13	15	16	19	22
Diletti <i>et al.</i> (1991)	noncentr. t	Owen's Q	4	5	7	NA	9	NA	12	NA	15	NA	19	NA
nQuery Advisor 7 (2007)	noncentr. t	AS 184	4	6	8	8	10	12	12	14	16	16	20	22
FARTSSIE 1.7 (2010)	noncentr. t	AS 243	4	5	7	8	9	11	12	13	15	16	19	22
EEG 2 01 (2009)	noncentr. t	AS 243	4	5	7	8	9	11	12	13	15	16	19	22
	brute force	ElMaestro	4	5	7	8	9	11	12	13	15	16	19	22
StudySize 2.0.1 (2006)	central t	?	NA	5	7	8	9	11	12	13	15	16	19	22
Hauschke <i>et al.</i> (1992)	approx. t		NA	NA	8	8	10	12	12	14	16	16	20	22
Chow & Wang (2001)	approx. t		NA	6	6	8	8	10	12	12	14	16	18	22
Kieser & Hauschke (1999) approx. <i>t</i>		2	NA	6	8	NA	10	12	14	NA	16	20	24
			CV%	, D										
original values	Method	Algorithm	22.5	24	4 25	26	27.5	28	30	32	34	36	38	40
PowerTOST 1.1-02 (2013)	exact	Owen's Q	24	2	6 28	30	34	34	40	44	50	54	60	66
Patterson & Jones (2006)	noncentr. t	AS 243	23	2	6 28	30	33	34	39	44	49	54	60	66
Diletti <i>et al.</i> (1991)	noncentr. t	Owen's Q	23	NA	A 28	NA	33	NA	39	NA	NA	NA	NA	NA
nQuery Advisor 7 (2007)	noncontr t	10404		_					4.0		50	54	60	66
	noncentr. t	AS 184	_ 24	2	6 28	30	34	34	40	44	50	94		
FARTSSIE 1.7 (2010)	noncentr. t	AS 184 AS 243	24 23	20 20	6 28 6 28	30 30 30	34 33	34 34	40 39	44 44	49	54	60	66
FARTSSIE 1.7 (2010)	noncentr. <i>t</i> noncentr. <i>t</i>	AS 184 AS 243 AS 243	24 23 23	20 20 20	6 28 6 28 6 28	30 30 30 30	34 33 33	34 34 34	40 39 39	44 44 44	49 49	54 54	60 60	66 66
FARTSSIE 1.7 (2010) EFG 2.01 (2009)	noncentr. <i>t</i> noncentr. <i>t</i> brute force	AS 184 AS 243 AS 243 ElMaestro	24 23 23 23	20 20 20 20	6 28 6 28 6 28 6 28 6 28	30 30 30 30 30 30	34 33 33 33 33	34 34 34 34	40 39 39 39	44 44 44 44	49 49 49 49	54 54 54	60 60 60	66 66 66
FARTSSIE 1.7 (2010) EFG 2.01 (2009) StudySize 2.0.1 (2006)	noncentr. t noncentr. t brute force central t	AS 184 AS 243 AS 243 ElMaestro ?	24 23 23 23 23 23	20 20 20 20 20 20	6 28 6 28 6 28 6 28 6 28 6 28	30 30 30 30 30 30 30 30 30	34 33 33 33 33 33	34 34 34 34 34	40 39 39 39 39	44 44 44 44 44	49 49 49 49 49	54 54 54 54 54	60 60 60 60	66 66 66 66
FARTSSIE 1.7 (2010) EFG 2.01 (2009) StudySize 2.0.1 (2006) Hauschke <i>et al.</i> (1992)	noncentr. t noncentr. t brute force central t approx. t	AS 184 AS 243 AS 243 ElMaestro ?	24 23 23 23 23 23 24	20 20 20 20 20 20 20	6 28 6 28 6 28 6 28 6 28 6 28 6 28 6 28 6 28 6 28 6 28 6 28	30 30 30 30 30 30 30 30 30 30 30 30	34 33 33 33 33 33 33 33 34	34 34 34 34 34 34 36	40 39 39 39 39 40	44 44 44 44 44 44 46	49 49 49 49 49 50	54 54 54 54 54 56	60 60 60 60 64	66 66 66 66 70
FARTSSIE 1.7 (2010) EFG 2.01 (2009) StudySize 2.0.1 (2006) Hauschke <i>et al.</i> (1992) Chow & Wang (2001)	noncentr. t noncentr. t honcentr. t brute force central t approx. t approx. t	AS 184 AS 243 AS 243 ElMaestro ?	24 23 23 23 23 23 24 24	20 20 20 20 20 20 20 20 20	5 28 6 28 6 28 6 28 6 28 6 28 6 28 6 28 6 28 6 28 6 28 6 28 6 28 6 28	30 30 30 30 30 30 30 30 30 30 30 30 30 3	34 33 33 33 33 33 34 34 34	34 34 34 34 34 36 34	40 39 39 39 39 40 38	44 44 44 44 44 46 44	49 49 49 49 50 50	54 54 54 54 54 56 56	60 60 60 60 64 62	66 66 66 70 68



Sample size tables

Diletti E, Hauschke D, and VW Steinijans

Sample size determination for bioequivalence assessment by means of confidence intervals Int J Clin Pharmacol Ther Toxicol 29(1), 1–8 (1991)

(α 0.05, ⊿ 0.2 [0.80 – 1.25], Power 80%						lpha 0.05, $arDelta$ 0.2 [0.80 – 1.25], Power 80%						,		$\alpha 0.05$	5, <u> </u>	.2 [0.8	<u> 30 – 1</u>	.25],	Powe	r 90%	
C\/0/			P	E (GN	1R, T/	R)			C\/0/			Ρ	E (GN	1R, T/	R)							
C v 70	0.85	0.90	0.95	1.00	1.05	1.10	1.15	1.20	C v %	0.85	0.90	0.95	1.00	1.05	1.10	1.15	1.20					
5.0	11	5	4	4	4	5	7	22	5.0	14	6	4	4	4	5	8	28					
7.5	21	7	5	5	5	7	12	44	7.5	28	9	6	5	6	8	16	60					
10.0	35	11	7	6	7	10	20	75	10.0	48	14	8	7	8	13	26	104					
12.5	54	16	9	8	9	14	30	117	12.5	74	21	11	9	11	18	40	161					
15.0	77	22	12	10	12	19	41	167	15.0	106	29	15	12	15	25	57	231					
17.5	103	29	15	13	15	25	56	226	17.5	142	39	20	15	19	34	75	312					
20.0	134	37	19	16	18	32	72	293	20.0	185	50	26	19	24	43	99	405					
22.5	168	46	23	19	23	39	90	368	22.5	232	63	31	23	30	54	124	509					
25.0	206	56	28	23	27	48	110	452	25.0	284	77	37	28	36	65	151	625					
27.5	247	67	33	27	33	57	132	543	27.5	342	92	44	34	43	78	181	751					
30.0	292	79	39	32	38	67	155	641	30.0	403	108	52	39	51	92	214	888					



Sample size tables

•Tóthfalusi L and L Endrényi

Sample Sizes for Designing Bioequivalene Studies for Highly Variable Drugs J Pharm Pharmaceut Sci 15(1), 73–84 (2011)

lpha 0.05, ABEL (EMA), partial repl., Power 80%						0%	α0.	05, R	SABE	E (FDA	A), pa	rtial re	pl., Po	ower 8	80%		
C\/%	PE (GMR, T/R)					PE (GMR, T/R)											
C V 70	0.85	0.90	0.95	1.00	1.05	1.10	1.15	1.20	C v 70	0.85	0.90	0.95	1.00	1.05	1.10	1.15	1.20
30	194	53	27	22	26	45	104	>201	30	145	45	24	21	24	39	82	>201
35	127	51	29	25	29	45	84	>201	35	74	37	24	22	25	34	54	109
40	90	44	29	27	30	42	68	139	40	60	33	24	22	24	31	47	104
45	77	40	29	27	29	37	57	124	45	59	31	23	22	24	29	43	116
50	75	40	30	28	30	37	53	133	50	66	30	24	22	23	28	41	133
55	81	42	32	30	32	40	56	172	55	80	30	24	22	24	28	44	172
60	88	46	36	33	36	44	63	>201	60	88	31	24	23	24	30	50	>201
65	99	53	40	37	40	50	71	>201	65	98	32	25	24	25	31	53	>201
70	109	58	45	41	45	56	80	>201	70	106	35	26	25	26	31	62	>201
75	136	67	50	46	50	62	89	>201	75	136	38	27	26	27	34	70	>201
80	144	72	54	51	55	68	97	>201	80	144	40	40	27	29	37	76	>201



Sample size tables

- •Never interpolate!
- •Use the most conservative cell entry (higher CV, PE away from 1)
 - Example: Sample size for CV 18%, PE 0.92, 80% power?

	PE (GMR, T/R)									
	0.90	0.95	1.00							
17.5	29	15	13							
20.0	37	19	16							



Round up to next even number (38)



Tables vs. calculations

- The penalty to be paid using tables might be high especially if uprounding has to be applied.
 - Sample sizes of the example: CV 18%, PE 0.92, 80% power
 - Table: n = 38
 - Approximations
 - Hauschke et al. 1992: n = 24
 - Chow and Wang 2001: n = 22
 - FARTSSIE.xls: n = 22

• Exact: n = 22



Tables vs. calculations

- If we planned the study in 38 subjects (tables) instead of the required 22 (exact) we gain a lot of power, but how much?
 - n = 22: power 80.55%
 - n = 38: power 95.56%
- If step sizes are too wide, calculations mandatory
 PowerTOST supports simulations for ABEL (EMA-method) and RSABE (FDA-method)



Tables vs. calculations

library(PowerTOST)
sampleN.scABEL(CV=0.40, details=F)

+++++ scaled (widened) ABEL ++++++ Sample size estimation

Study design: 2x3x3
log-transformed data (multiplicative
model)
1e+05 studies simulated.

```
alpha = 0.05, target power = 0.8
CVw(T) = 0.4; CVw(R) = 0.4
Null (true) ratio = 0.95
ABE limits/PE constraints = 0.8..1.25
Regulatory settings: EMA
- CVswitch = 0.3, cap on ABEL
```

if CV > 0.5
- Regulatory constant = 0.76

Sample size n power 30 0.827170

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library(PowerTOST)
sampleN.RSABE(CV=0.40, details=F)

++++ Reference scaled ABE crit. ++++ Sample size estimation

Study design: 2x3x3
log-transformed data (multiplicative
model)
1e+05 studies simulated.

```
alpha = 0.05, target power = 0.8
CVw(T) = 0.4; CVw(R) = 0.4
Null (true) ratio = 0.95
ABE limits/PE constraints = 0.8...1.25
Regulatory settings: FDA
```

```
Sample size
n power
24 0.808640
```



•ICH E9 (1998)

Section 3.5 Sample Size, paragraph 3

- The method by which the sample size is calculated should be given in the protocol [...]. The basis of these estimates should also be given.
- It is important to investigate the sensitivity of the sample size estimate to a variety of deviations from these assumptions and this may be facilitated by providing a range of sample sizes appropriate for a reasonable range of deviations from assumptions.
- In confirmatory trials, assumptions should normally be based on published data or on the results of earlier trials.



•Example

20% n=2

nQuery Advisor: $\sigma_{w} = \sqrt{\ln(CV_{intra}^{2} + 1)}; \sqrt{\ln(0.2^{2} + 1)} = 0.198042$

🔯 nQuery Advisor - [MTE2co-1.nqa]						
<u> R</u> ile Edit <u>V</u> iew <u>O</u> ptions <u>A</u> ssistants <u>R</u> ai	ndomize <u>P</u> lot	<u>W</u> indow <u>H</u>	elp			
		🐨 🔳 St	K T			
t-tests (TOST) of equivalence in ratio of mea	ns for crosso	ver design (n	atural log scale)	l	_	
	90% power	25% CV	4 drop outs	25% CV + d.o.	PE 90%	worst case
Test significance levels, α (one-sided)	0.050	0.050	0.050	0.050	0.050	0.050
Lower equivalence limit for μ_T / μ_S, Δ_L	0.800	0.800	0.800	0.800	0.800	0.800
Upper equivalence limit for μ_T / μ_S,Δ_U	1.250	1.250	1.250	1.250	1.250	1.250
Expected ratio, μ_T / μ_S	0.950	0.950	0.950	0.950	0.900	0.900
Crossover ANOVA, sqrt(MSE) (In scale)	0.198042	0.246221	0.198042	0.246221	0.198042	0.246221
SD differences, ơ _d (In scale)	0.280074	0.348209	0.280074	0.348209	0.280074	0.348209
Power (%)	90.00	77.60	86.88	69.53	66.94	45.09
n per sequence group	13	13	11	11	13	11
1						
25% CV:	0% CV, 4 d ower 90%	rop outs: $\rightarrow 87\%$	25% C	V, 4 drop outs	20% powe	CV, PE 90%: er 90% → 67
			power	<u>'90% → 70%</u>		



•Example *PowerTOST*, function *sampleN.TOST*

+++++++++ Equivalence test - TOST +++++++++ Sample size estimation

```
Study design: 2x2 crossover
log-transformed data (multiplicative model)
alpha = 0.05, target power = 0.9
BE margins = 0.8 ... 1.25
Null (true) ratio = 0.95, CV = 0.2
Sample size
n power
26 0.917633
```



•To estimate Power for a given sample size, use function *power*.*TOST*

library(PowerTOST)
power.TOST(alpha=0.05, theta0=0.95, CV=0.25, n=26, design="2x2")
[1] 0.7760553

```
power.TOST(alpha=0.05, theta0=0.95, Cv=0.20, n=22, design="2x2")
[1] 0.8688866
```

```
power.TOST(alpha=0.05, theta0=0.95, CV=0.25, n=22, design="2x2")
[1] 0.6953401
```

```
power.TOST(alpha=0.05, theta0=0.90, CV=0.20, n=26, design="2x2")
[1] 0.6694514
```

```
power.TOST(alpha=0.05, theta0=0.90, CV=0.25, n=22, design="2x2")
[1] 0.4509864
```



•Must be done *before* the study (a priori)

•The Myth of retrospective (*a posteriori*) Power...

High values do not further support the claim of already demonstrated bioequivalence.

Low values do not invalidate a bioequivalent formulation.
 Further reader:

RV Lenth (2000) JM Hoenig and DM Heisey (2001) P Bacchetti (2010)



Thank You! Sample Size Estimation for BE Studies Open Questions?



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To bear in Remembrance...

Power. That which statisticians are always calculating but never have.

Power: That which is wielded by the priesthood of clinical trials, the statisticians, and a stick which they use to beta their colleagues.



Power Calculation – A guess masquerading as mathematics.

Stephen Senn



You should treat as many patients as possible with the new drugs while they still have the power to heal. Armand Trousseau



The Myth of Power

There is simple intuition behind results like these: If my car made it to the top of the hill, then it is powerful enough to climb that hill; if it didn't, then it obviously isn't powerful enough. Retrospective power is an obvious answer to a rather uninteresting question. A more meaningful question is to ask whether the car is powerful enough to climb a particular hill never climbed before; or whether a different car can climb that new hill. Such questions are prospective, not retrospective.

The fact that retrospective power adds no new information is harmless in its own right. However, in typical practice, it is used to exag-



gerate the validity of a significant result ("not only is it significant, but the test is really powerful!"), or to make excuses for a nonsignificant one ("well, P is .38, but that's only because the test isn't very powerful"). The latter case is like blaming the messenger.

RV Lenth

Two Sample-Size Practices that I don't recommend http://www.math.uiowa.edu/~rlenth/Power/2badHabits.pdf



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•Collection of links to global documents <u>http://bebac.at/Guidelines.htm</u>

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