

# Noncompartmental Analysis, Statistical Evaluation

# Noncompartmental Analysis

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- NCA a.k.a. SHAM (Shape, Height, Area, Moments)
  - PK metrics (plasma)
    - Single dose
      - Extent of Absorption (WHO, EEA, ...), Total Exposure (USA):  $AUC$  (Area Under the Curve)
        - » In most jurisdictions the PK metric for BE is  $AUC_{0-t}$ , where  $t$  is the last time point with a quantifiable concentration
        - » WHO, EEA: For IR products with a long half life  $AUC_{0-72}$  is sufficient
        - » USA and EEA (controlled release products only): additionally  $AUC_{0-\infty}$
      - Rate of Absorption (WHO, EEA, ...), Peak Exposure (USA):  $C_{max}$
      - $t_{max}$  (Russia, Eurasian Economic Area, ...)
      - Rarely relevant
        - »  $t_{75\%}$ , POT-25 (Plateau time or peak occupancy time; time span where  $C(t) \geq 75\% C_{max}$ : Russia for modified release products)
        - »  $MRT$  (Mean of Residence Times)
        - » Therapeutic Occupancy Time (time span where  $C(t) \geq$  some given limit, e.g., the MIC)

# Noncompartmental Analysis

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- Multiple dose
  - Extent of Absorption (WHO, EEA, ...), Total Exposure (USA):  
 $AUC_{0-\tau}$  (AUC covering the dosing interval  $\tau$ )  
If chronopharmacological variation and more than o.a.d. regimen:  
 $AUC_{0-24}$   
No extrapolation of  $AUC$  in any case
  - Rate of Absorption (WHO, EEA, ...), Peak Exposure (USA):  
 $C_{max,ss}$
  - Minimum concentration  
 $C_{min,ss}$  (lowest observed concentration within the profile; originators)  
 $C_{\tau,ss}$  (concentration at the end of the dosing interval; generics)
  - $PTF$  (Peak-to-Trough Fluctuation)  
 $(C_{max,ss} - C_{min,ss}) / C_{av,ss}$ , where  $C_{av,ss} = AUC_{0-\tau} / \tau$
  - Mentioned in some GLs but practically obsolete due to its extreme variability  
 $Swing = (C_{max,ss} - C_{min,ss}) / C_{min,ss}$

# Noncompartmental Analysis

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- PK metrics obtained by NCA depend much more on the sampling schedule than PK parameters estimated with a PK model
  - Examples
    - It is unlikely that one is able to ‘catch’ the true  $C_{max}/t_{max}$  in every subject
      - Hence, frequent sampling around  $t_{max}$  mandatory
    - To obtain a reliable estimate of the apparent elimination  $\lambda_z$ , *at least* three samples required
  - However, contrary to PK modeling NCA is independent from software
    - Paper, pencil, brain...

# PK model | AUC

- AUC is the integral of the concentration-time curve
  - One compartment, extravascular dose, no lag-time

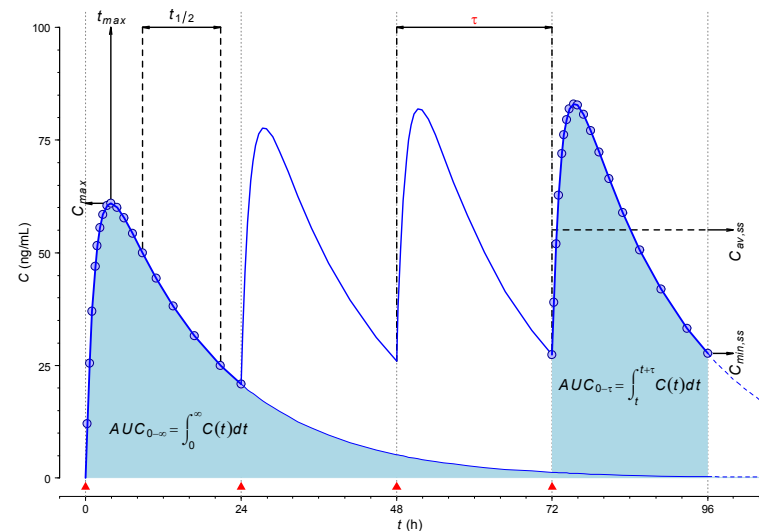
$$C(t) = \frac{f \cdot D}{V} \frac{k_a}{k_a - k_e} \left( e^{-k_e \cdot t} - e^{-k_a \cdot t} \right)$$

$$AUC_{0-\infty} = \int_0^{\infty} C(t) dt$$

$$= \frac{f \cdot D}{V} \frac{k_a}{k_a - k_e} \left( \frac{1}{k_e} - \frac{1}{k_a} \right)$$

$$= \frac{f \cdot D}{V \cdot k_e}$$

$$= \frac{f \cdot D}{CL}$$



Superposition Principle of linear PK  
 $AUC_{0-T} \approx AUC_{0-\infty}$

# NCA | *AUC*

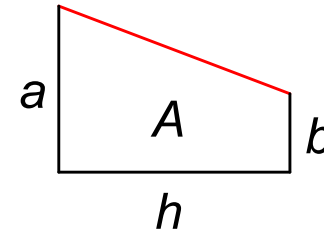
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- In NCA numeric approximation of the integral is required
  - Linear trapezoidal method
  - Linear-up / logarithmic-down trapezoidal method
  - Of academic interest
    - Cubic splines
    - Lagrange polynomials
    - Simpson's rule

# AUC | linear trapezoidal method

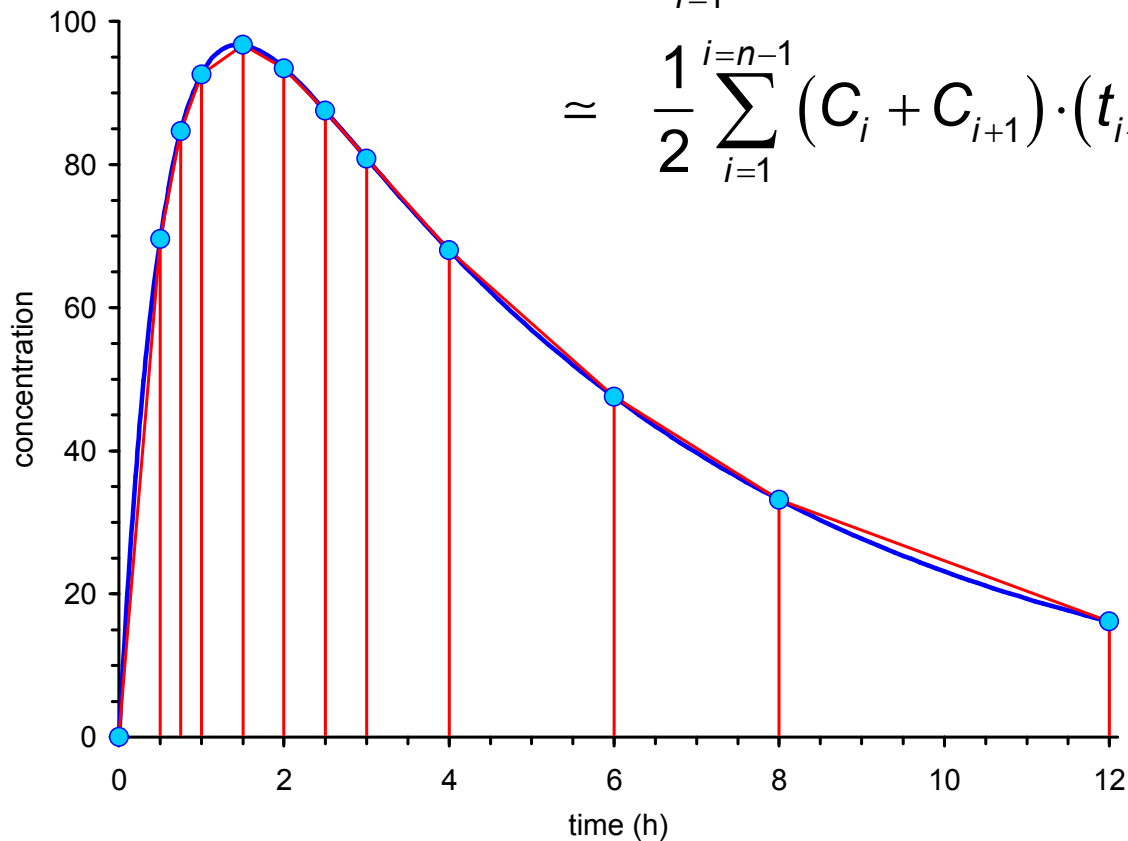
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- Linear interpolation between data points
- Sections are represented by trapezoids
- Sides  $a$ ,  $b$  are two neighbouring concentrations
- $h$  is the time interval
- Area of one trapezoid  $A = \frac{a+b}{2}h$



# AUC | linear trapezoidal method

- Total Area  $AUC_{0-t_n} \approx \sum_{i=1}^{i=n-1} \frac{C_i + C_{i+1}}{2} (t_{i+1} - t_i)$   
 $\approx \frac{1}{2} \sum_{i=1}^{i=n-1} (C_i + C_{i+1}) \cdot (t_{i+1} - t_i)$

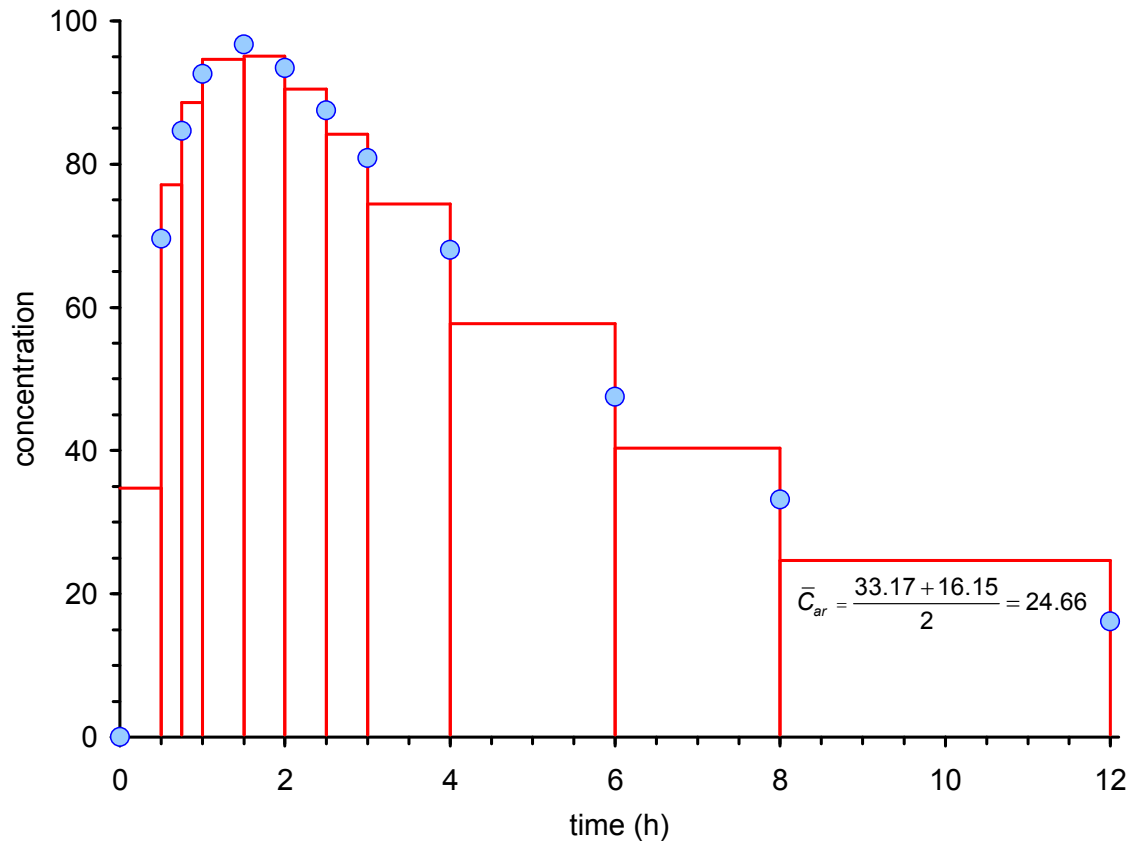




# AUC | linear trapezoidal method

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*arithmetic means of neighbouring concentrations*



# AUC | linear trapezoidal method

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- Positive bias
  - Overestimates *AUC* in both the absorption and distribution / elimination phases
- Originated in the dark ages
  - when profiles were plotted on paper, cut out, weighed on an analytical scale, and compared to the paper-weight of known area (e.g., A4 of 80 g/m<sup>2</sup>: 4.9896 g / 623.7 cm<sup>2</sup>)
- Should have been thrown into the scientific waste-can with the invention of pocket calculators decades ago
- In general elimination follows an exponential decrease

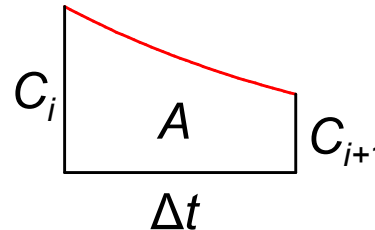
$$C(t) = \frac{f \cdot D}{V} \frac{k_a}{k_a - k_e} (e^{-k_e \cdot t} - e^{-k_a \cdot t})$$

# AUC | lin-up / log-down trapezoidal method

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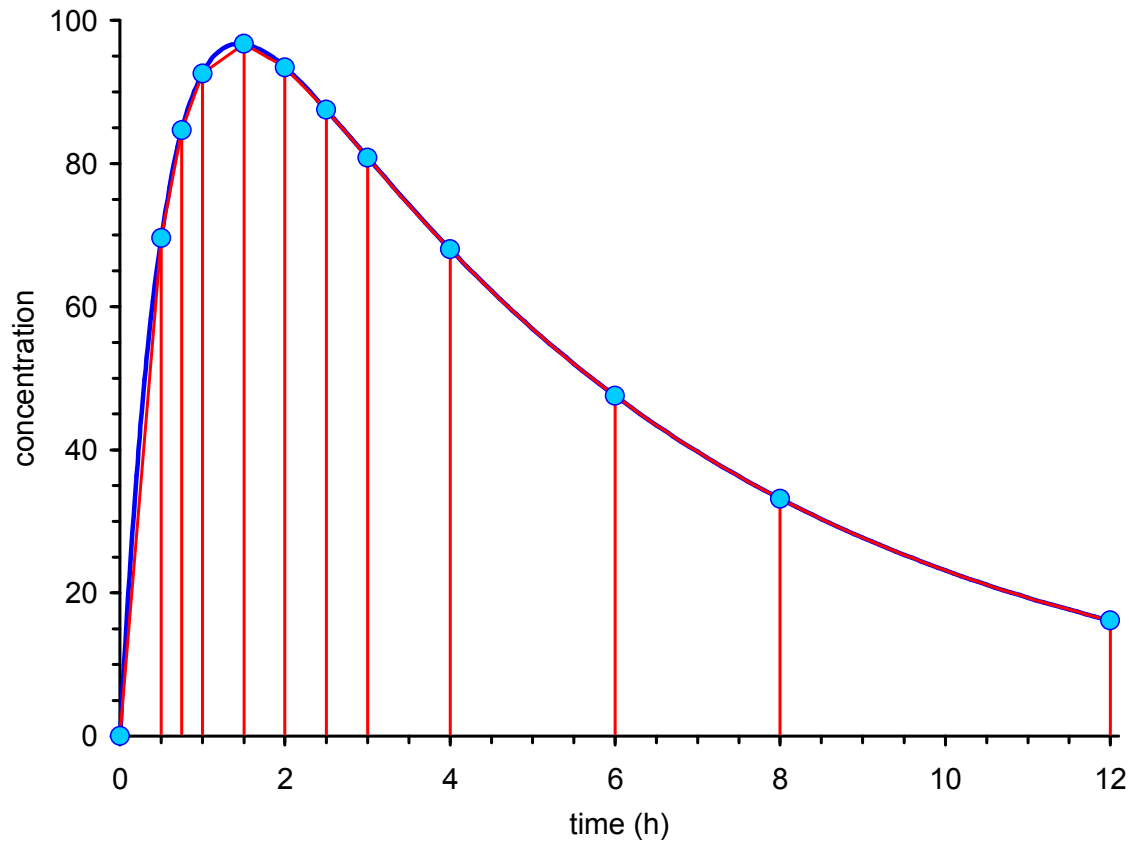
- Much better alternative:  
Linear-up / logarithmic-down trapezoidal method
- Sections with *increasing or equal* concentrations ( $C_{i+1} \geq C_i$ ) calculated by the linear trapezoidal method
- Sections with *decreasing* concentrations ( $C_{i+1} < C_i$ ) calculated by the logarithmic-linear trapezoidal method, *i.e.*,

$$AUC_{t_i-t_{i+1}} \simeq \frac{C_{i+1} - C_i}{\ln \frac{C_{i+1}}{C_i}} (t_{i+1} - t_i)$$



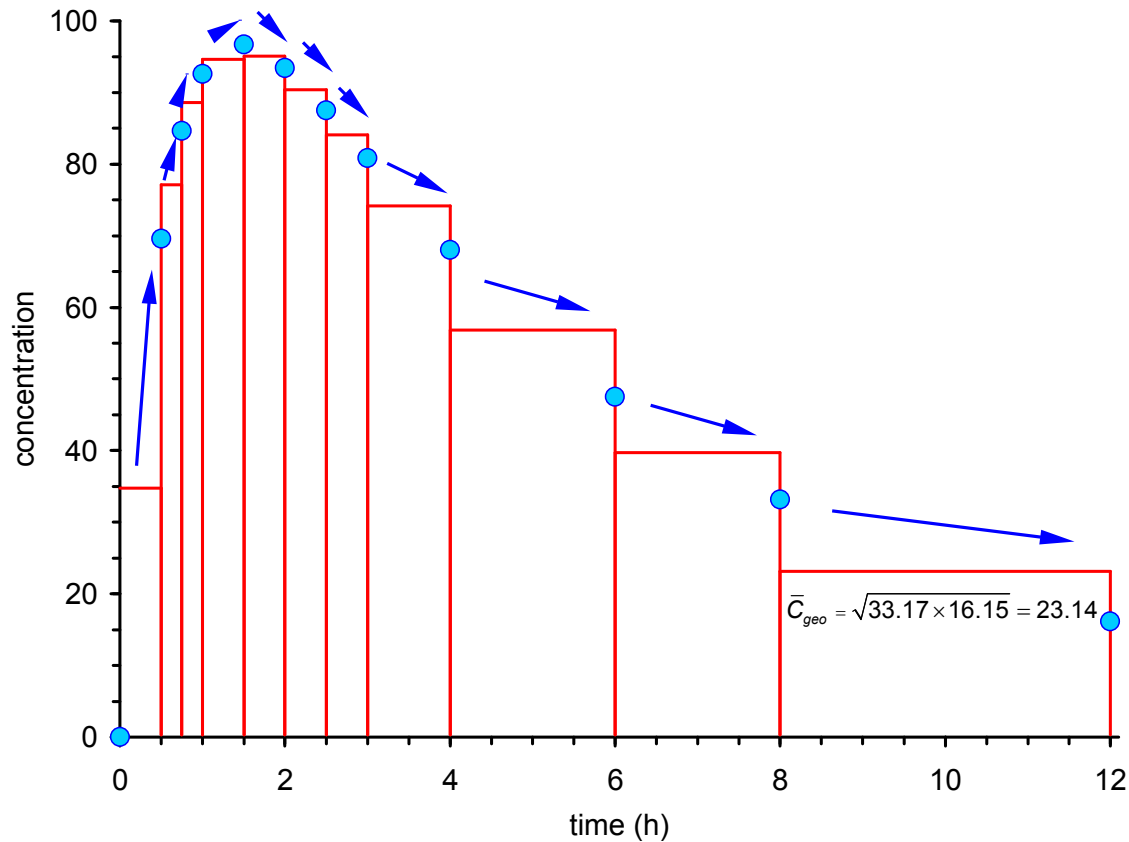
# AUC | lin-up / log-down trapezoidal method

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# AUC | lin-up / log-down trapezoidal method

*arithmetic / geometric means of neighbouring concentrations*



# AUC | lin-up / log-down trapezoidal method

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- Avoids positive bias in distribution / elimination phases
- Suitable for both i.v. and e.v. administrations
- Suitable for multiphasic profiles
  - Secondary peaks due to enterohepatic recycling
  - Pulsatile release products
  - If *AUC* of more than one profile has to be calculated (e.g., two doses with  $\tau$  12 h and  $AUC_{0-24}$  is required due to circadian variation in PK)
- Implemented in standard PK software for decades
- Only exception where the method performs worse than the linear trapezoidal
  - Drugs following Michaelis-Menten PK (e.g., alcohol)

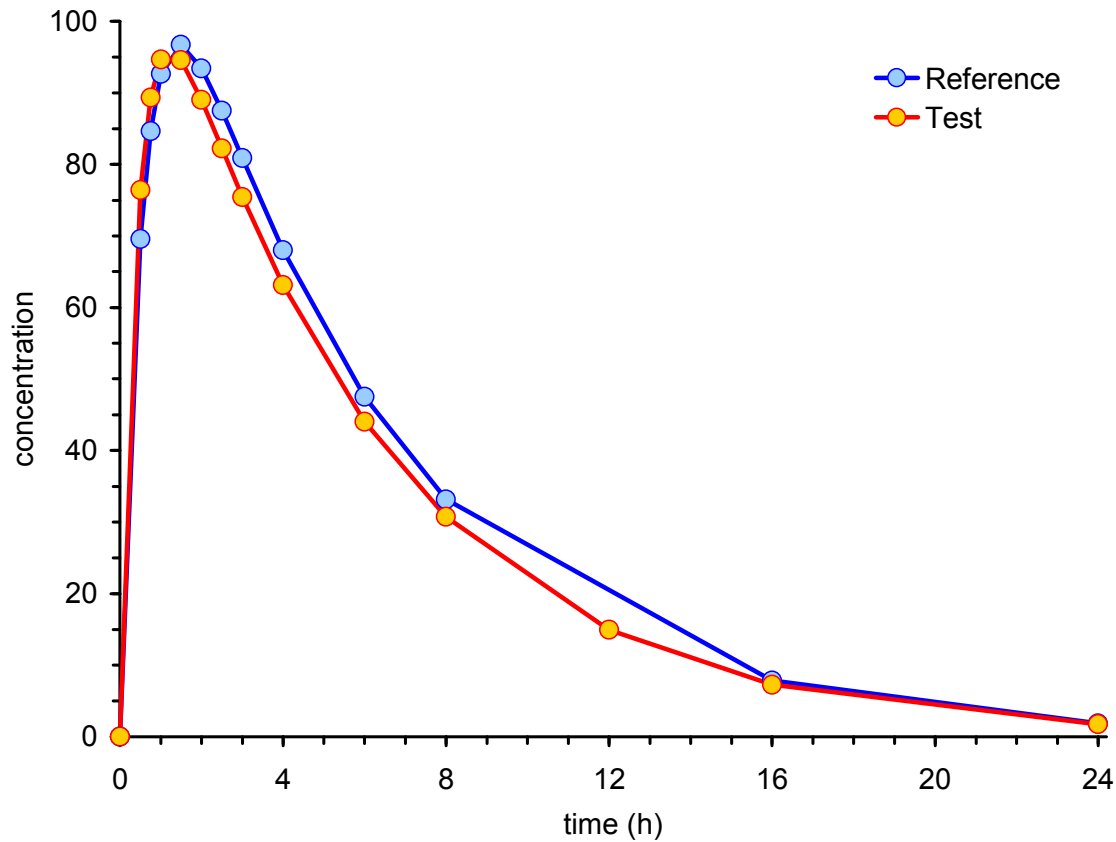
# $AUC_{0-t}$ | Problem 1

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- Recap: In most jurisdictions the PK metric for BE is  $AUC_{0-t}$ , where  $t$  is the last time point with a quantifiable concentration
- Ideally we are able to calculate  $AUC_{0-t}$ 
  - for all treatments
  - in all subjects
- What if
  - a sample was missing (e.g., vial broken in centrifugation)?
- Example
  - True T/R-ratio 95%, 12 h sample (R) missing
  - Comparison of linear and lin-up / log-down trapezoidal methods

# $AUC_{0-t}$ | Problem

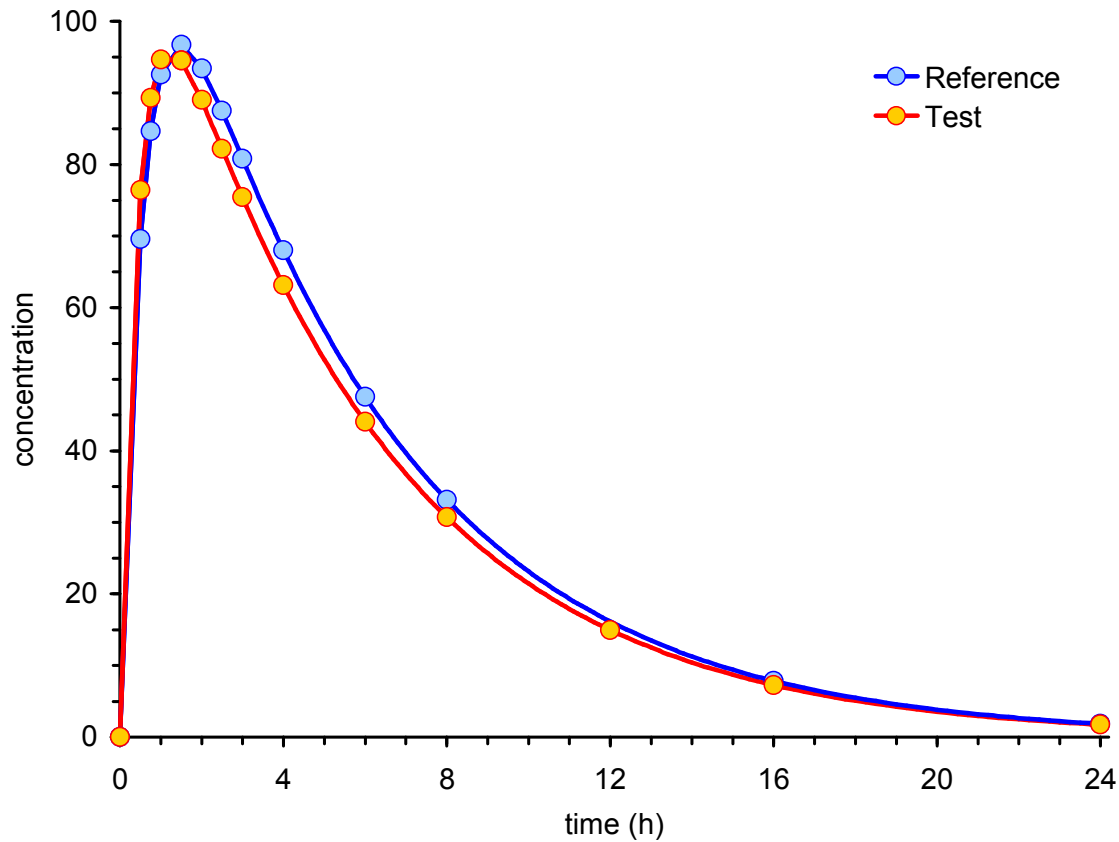
$AUC^\infty$  (R) 725,  $AUC^\infty$  (T) 671, T/R 92.5%, bias -2.60%





# $AUC_{0-t}$ | Solution

$AUC^\infty$  (R) 694,  $AUC^\infty$  (T) 658, T/R 94.9%, bias -0.15%



# $C_{max}$ | Problem & Solutions

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- What if
  - samples in the area of  $t_{max}$  are missing?
- Exclude the subject from the comparison of  $C_{max}$ 
  - Power depends on the  $CV$  (coefficient of variation), the  $GMR$  (geometric mean ratio), and  $n$  (sample size) where the rank order of their influence on power is
$$GMR \gg CV > n$$
  - Power will be compromised but to a much lesser degree than many people expect

- Recap: To obtain a reliable estimate of the apparent elimination  $\lambda_z$ , *at least* three samples required
  - The automatic algorithm based on maximizing  $R^2_{adj}$  is known to be ‘greedy’ (*i.e.*, reaches for too early time points) and
    - has difficulties with ‘flat’ profiles (*e.g.*, ill-defined  $C_{max}$  of controlled release products) and
    - regularly fails completely for multiphasic release products
  - Alternative: TTT method \*
    - Implemented in the open source package *bear* for *R*
  - Visual inspection of fits by a pharmacokineticist (with optional correction) is mandatory in all methods

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\* Scheerans C, Derendorf H, Kloft C. *Proposal for a Standardised Identification of the Mono-Exponential Terminal Phase for Orally Administered Drugs*. Biopharm Drug Dispos. 2008;29(3):145–57. doi:10.1002/bdd.596.

# Statistical Evaluation

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- Parallel Designs

- One group is treated with the test formulation and another group with the reference
- Quite common that – due to dropouts – the data set of eligible subjects is imbalanced, *i.e.*,  $n_1 \neq n_2$
- Equal variances should never be assumed (FDA 2001)
  - Treatment effect might be biased and
  - patient's risk inflated
  - In some software (*e.g.*, Kinetica, ThothPro) either wrong calculation or not possible at all

Subj.	Group 1 (T)	Group 2 (R)
1-13	100	110
2-14	103	113
3-15	80	96
4-16	110	90
5-17	78	111
6-18	87	68
7-19	116	111
8-20	99	93
9-21	122	93
10-22	82	82
11-23	68	96
12-24	dropout	137
<i>n</i>	11	12
mean	95	100
<i>s</i> <sup>2</sup>	298	314
<i>s</i>	17.3	17.7

# Statistical Evaluation

Subj.	Group 1 (T)	ln (T)	Group 2 (R)	ln (R)
1-13	100	4.605	110	4.700
2-14	103	4.635	113	4.727
3-15	80	4.382	96	4.564
4-16	110	4.700	90	4.500
5-17	78	4.357	111	4.710
6-18	87	4.466	68	4.220
7-19	116	4.754	111	4.710
8-20	99	4.595	93	4.533
9-21	122	4.804	93	4.533
10-22	82	4.407	82	4.407
11-23	68	4.220	96	4.564
12-24	dropout	–	137	4.920
<i>n</i>	11	11	12	12
mean	95	4.539	100	4.591
<i>s</i> <sup>2</sup>	298	0.03418	314	0.03231
<i>s</i>	17.3	0.1849	17.7	0.1798

Assuming equal variances

$$\nu = n_1 + n_2 - 2 = 21$$

$$t_{1-\alpha,21} = 1.7207$$

90% CI: 83.28% – 108.20%

Adjusting for unequal variances by Satterthwaite's degrees of freedom

$$\nu = \frac{(s_1^2/n_1 + s_2^2/n_2)^2}{\frac{(s_1^2/n_1)^2}{n_1 - 1} + \frac{(s_2^2/n_2)^2}{n_2 - 1}} = 20.705$$

$$t_{1-\alpha,20.705} = 1.7219$$

90% CI: 83.26% – 108.23%

Minor difference in the CIs but only little imbalance in the data and variances quite similar. However, the simple *t*-test is always liberal, *i.e.*, compromises the patient's risk.

# Statistical Evaluation

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- Crossover Designs ( $2 \times 2 \times 2$ )
  - Every subject is treated with both the test and the reference formulation
  - Subjects randomized to two sequences TR and RT
  - Treatment periods separated by washout
  - Potential period effects are accounted for in the analysis (mean out)
  - Evaluation by
    - Analysis of Variance (ANOVA) – WHO, EMA, ...
    - Linear mixed effects model – FDA, Health Canada
    - Results are identical for balanced datasets (equal number of subjects in both sequences) and differ only slightly for imbalanced ones

# Statistical Evaluation

- Crossover Designs ( $2 \times 2 \times 2$ ) – Example

subject	T	R
1	28.39	35.44
2	39.86	49.42
3	32.75	36.78
4	33.36	33.40
5	34.97	34.81
6	24.29	24.65
7	28.61	31.77
8	45.44	45.54
9	59.49	65.29
10	27.87	28.23
11	24.26	25.71
12	42.30	37.01

sequence RT			sequence TR		
subject	P I	P II	subject	P I	P II
2	39.86	49.42	1	28.39	35.44
3	32.75	36.78	4	33.36	33.40
5	34.97	34.81	6	24.29	24.65
8	45.44	45.54	7	28.61	31.77
10	27.87	28.23	9	59.49	65.29
11	24.26	25.71	12	42.30	37.01

Ordered by treatment sequences (RT|TR)

ANOVA on log-transformed data →

# Statistical Evaluation

- Crossover Designs (2×2×2) – Example cont'd

Sequence	Period 1		Period 2		Sequence mean	
1	1R = $X_{\cdot 11}$	3.5103	1T = $X_{\cdot 21}$	3.5768	$X_{\cdot \cdot 1}$	3.5436
2	2T = $X_{\cdot 12}$	3.5380	2R = $X_{\cdot 22}$	3.5883	$X_{\cdot \cdot 2}$	3.5631
Period mean	$X_{\cdot 1 \cdot}$	3.5241	$X_{\cdot 2 \cdot}$	3.5826	$X_{\cdot \cdot \cdot}$	3.5533
RT = $n_1 = 6$						
TR = $n_2 = 6$ $1/n_1 + 1/n_2$ 0.3333						
<b>balanced</b> $n = 12$ $1/n$ 0.0833 $n_1 + n_2 - 2$ 10						
<b>Analysis of Variance</b>						
Source of variation	df	SS	MS	F	P-value	CV
<i>Inter-subjects</i>						
Carry-over	1	0.00230	0.00230	0.0144	0.90679	
Residuals	10	1.59435	0.15943	29.4312	4.32E-6	28.29%
<i>Intra-subjects</i>						
Direct drug	1	0.00040	0.00040	0.0733	0.79210	
Period	1	0.02050	0.02050	3.7844	0.08036	
Residuals	10	0.05417	0.00542			7.37%
Total	23	1.67172				

$\delta_{ML}$  **1.0082** MLE (maximum likelihood estimator) of Delta-ML

$X_R$  3.5493 LS (least squares mean for the reference formulation)     $\exp(X_R)$  34.79

$X_T$  3.5574 LS (least squares mean for the test formulation)     $\exp(X_T)$  35.07



# Statistical Evaluation

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- Crossover Designs (2×2×2) – Example cont'd

### Classical (Shortest) Confidence Interval

± x rule: **20** [ 100 - x; 1 / (100 - x) ]

$\theta_L$  -0.2231                       $\theta_U$  +0.2231                       $\alpha$  0.0500  $p=1-2\cdot\alpha$  0.9000

$\delta_L$      **80%**                               $\delta_U$      **125%**                       $t_{2,\alpha,df}$  1.8125

$L_1$  -0.0463                               $U_1$      0.0626 *difference within Theta-L AND Theta-U; bioequivalent*

$L_2$  **95.47%**                               $U_2$      **106.46%** *difference within Delta-L AND Delta-U; bioequivalent*

$\delta_{ML}$   $\hat{\leftarrow}$  **100.82%**  $\hat{\rightarrow}$  *MLE; maximum likelihood estimator*

$\delta_{MVUE}$      100.77%                      *MVUE; minimum variance unbiased estimator*

$\delta_{RM}$      100.98%                      *RM; ratio of formulation means*

$\delta_{MIR}$      101.44%                      *MIR; mean of individual subject ratios*

# Statistical Evaluation

- Interpreting ANOVA Tables – Example cont'd

Analysis of Variance						
Source of variation	df	SS	MS	F	P-value	CV
<i>Inter-subjects</i>						
Carry-over	1	0.00230	0.00230	0.0144	0.90679	
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Residuals	10	0.05417	0.00542			7.37%
Total	23	1.67172				

Should not be tested:  
Design – washout!

Not surprising:  
different subjects!

Irrelevant: Significant value  
would only mean that 100% is  
not included in the 90% CI.

Irrelevant; both formulations  
affected to the same degree

$$CV_{within} = 100\sqrt{e^{MS_w} - 1}$$

# Statistical Evaluation

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- Statistical significant  $\neq$  clinically relevant
  - For any given T/R-ratio and variability one will get a significant treatment effect (in the ANOVA  $p < 0.05$ ) if the sample size is only large enough
    - The confidence interval narrows with  $\sqrt{N}$ , *i.e.*, if one uses a four times larger sample size, the CI will be  $\sim$ half as wide
    - If the CI does not include 100% any more, treatments will *statistically significant* differ
    - However, if the 90% CI is within the BE-limits, this difference is *clinically not relevant*

# Statistical Evaluation

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- General Procedure (all Designs)
  - Based on the design set up a statistical model
  - Log-transform the PK metrics of interest
  - Calculate for T and R
    - Balance sequences: Geometric mean
    - Imbalanced sequences: Adjusted mean (a.k.a. least squares mean)
  - Calculate the ratio of means
  - Calculate the 90% confidence interval (CI) around the ratio
  - The *width* of the CI depends on the variability observed in the study
  - The *location* of the CI depends on the observed test/reference-ratio

# Statistical Evaluation

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- BE Assessment (all Designs)
  - Decision rules based on the CI and pre-specified BE-limits
    - CI entirely outside the BE-limits →  
**Bioinequivalence proven**
    - CI overlaps the BE-limits (lies not entirely within the limits) →  
**Bioequivalence not proven** (indecisive)
    - CI entirely inside the BE-limits →  
**Bioequivalence proven**
  - Methods for reference-scaling
    - The BE-limits depend on the  $CV_{wR}$  observed in study
    - Only the method pre-specified in the protocol

# Statistical Evaluation

