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Study Designs

The more 'sophisticated' a design is, the more information can be extracted.

Hierarchy of designs:

```
Full replicate (RTRT | TRTR or RTR | TRT) →
Partial replicate (RRT | RTR | TRR) →
2×2×2 crossover (RT | TR) →
Parallel (R | T)
```

Variances which can be estimated:

Parallel: total variance (pooled of between + within subjects)

2×2×2 crossover: + between, within subjects *→*

Partial replicate: + within subjects (of R)

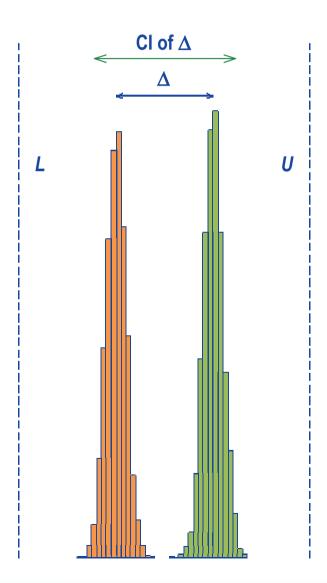
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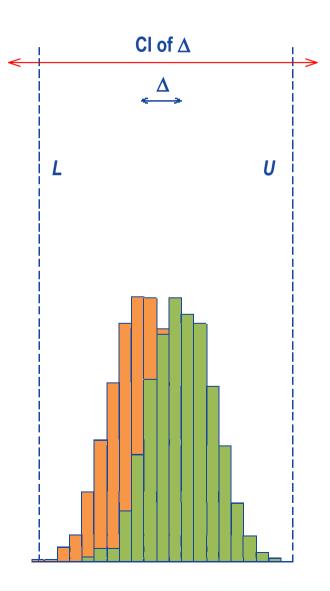
Full replicate: + within subjects (of R and T) *→*





Highly Variable Drugs / Drug Products





Counterintuitive concept of BE:

Two formulations with a large difference in means are declared bioequivalent if variances are low, but not BE – even if the difference is quite small – due to high variability.

Modified from Tothfálusi et al. (2009), Fig. 1



It may be almost impossible to demonstrate BE with a reasonable sample size.

- Reference-scaling (*i.e.*, widening the acceptance range based of the variability of the reference) in 2010 introduced by the FDA and EMA and in 2016 by Health Canada.
 - Requires a replicate design, where at least the reference product is administered twice.
 - Smaller sample sizes compared to a standard 2×2×2 design but outweighed by increased number of periods.
 - Similar total number of individual treatments.
 - Any replicate design can be evaluated for 'classical' (unscaled) Average Bioequivalence (ABE) as well. Switching CV_{wR} 30%:
 - FDA: AUC and C_{max}
 - EMA: C_{max} ; MR products additionally: $C_{ss,min}$, $C_{ss,r}$, partial AUCs
 - Health Canada: AUC





Models (in log-scale).

- ABE Model:
 - A difference \triangle of ≤20% is considered to be clinically not relevant.
 - The limits [L, U] of the acceptance range are fixed to $log(1 \Delta) = log((1 \Delta)^{-1})$ or $L \sim -0.2231$ and $U \sim +0.2231$.
 - The consumer risk is fixed with 0.05. BE is concluded if the $100(1 2\alpha)$ confidence interval lies entirely within the acceptance range.

$$-\theta_{A} \leq \mu_{T} - \mu_{R} \leq +\theta_{A}$$

- SABEL Model:
 - Switching condition θ_S is derived from the regulatory standardized variation σ_0 (proportionality between acceptance limits in log-scale and σ_{wR} in the highly variable region).

$$-\theta_{S} \leq \frac{\mu_{T} - \mu_{R}}{\sigma_{WR}} \leq +\theta_{S}$$



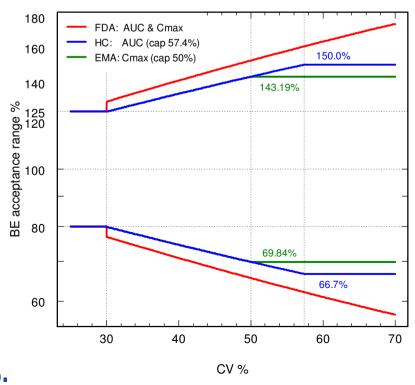


Regulatory Approaches.

• Bioequivalence limits derived from $\sigma_{\!_{0}}$ and $\sigma_{\!_{wR}}$

$$\theta_{S} = \frac{\log(1.25)}{\sigma_{0}}, [L,U] = e^{\pm\theta_{S}\cdot\sigma_{WR}}$$

- FDA
 - Scaling σ_{wR} 0.25 (θ_{S} 0.893) but applicable at $CV_{wR} \ge 30\%$.
 - Discontinuity at CV_{wR} 30%.
- EMA
 - Scaling σ_0 0.2936 (θ_S 0.760).
 - Upper cap at CV_{wR} 50%.
- Health Canada
 - Like EMA but upper cap at CV_{wR} 57.4%.





The EMA's Approach.

- Average Bioequivalence with Expanding Limits ABEL (crippled from Endrényi and Tóthfalusi 2009).
 - Justification that the widened acceptance range is clinically not relevant (important – different to the FDA).
 - Assumes identical variances of T and R [sic] like in a 2×2×2.
 - All fixed effects model according to the Q&A-document preferred.
 - Mixed-effects model (allowing for unequival variances) is 'not compatible with CHMP guideline'...
 - Scaling limited at a maximum of CV_{wR} 50% (i.e., to 69.84 143.19%).
 - GMR within 0.8000 1.2500.
 - Demonstration that $CV_{wR} > 30\%$ is not caused by outliers (box plots of studentized intra-subject residuals?)...
 - — ≥12 subjects in sequence RTR of the 3-period full replicate design.

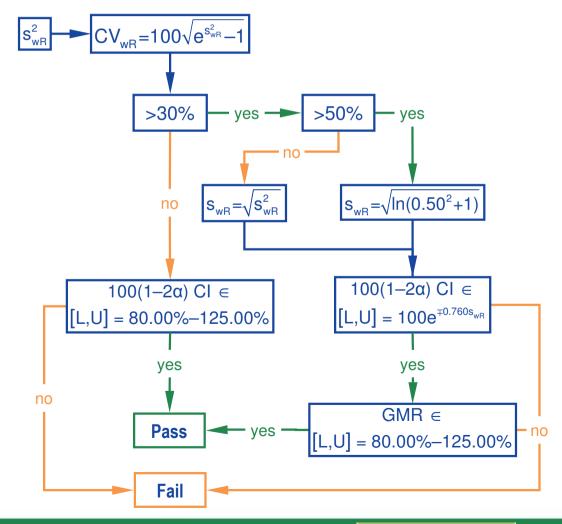




The EMA's Approach.

- Decision Scheme.
 - The Null Hypothesis is specified in the face of the data.
 - Acceptance limits themselves become random variables.
 - Type I Error (consumer risk) might be inflated.







Assessing the Type I Error (TIE).

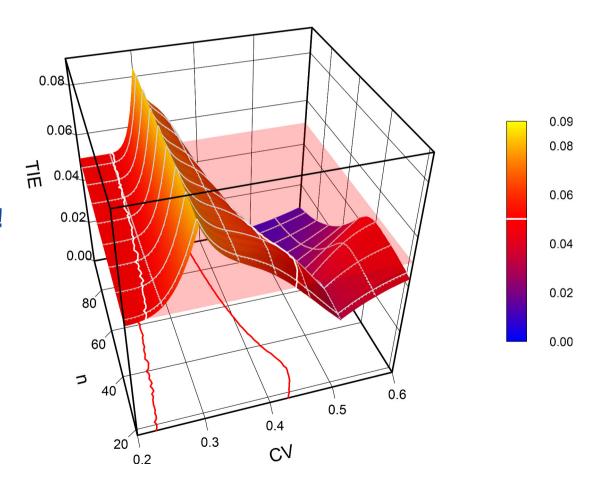
- TIE = falsely concluding BE at the limits of the acceptance range.
- Due to the decision scheme direct calculation of the TIE at the scaled limits is not possible;
 - \rightarrow extensive simulations required (10⁶ BE studies mandatory).
- Inflation of the TIE suspected. (Chow et al. 2002, Willavazie/Morgenthien 2006, Chow/Liu 2009, Patterson/Jones 2012).
- Confirmed.
 - EMA's ABEL
 (Tóthfalusi/Endrényi 2009, BEBA-Forum 2013, Wonnemann et al. 2015, Muñoz et al. 2016, Labes/Schütz 2016).
 - FDA's RSABE
 (Tóthfalusi/Endrényi 2009, BEBA-Forum 2013, Muñoz et al. 2016).





Example for ABEL

- RTRT | TRTR
 sample size 18 96
 CV_{wR} 20% 60%
 - TIE_{max} 0.0837.
 - Relative increase of the consumer risk 67%!







What is going on here?

SABE is stated in model parameters ...

$$-\theta_{S} \leq \frac{\mu_{T} - \mu_{R}}{\sigma_{WR}} \leq +\theta_{S}$$

- ... which are unknown.
- Only their estimates (GMR, s_{wR}) are accessible in the actual study.
- At CV_{wR} 30% the decision to scale will be wrong in ~50% of cases.
- If moving away from 30% the chances of a wrong decision decrease and hence, the TIE.
- At high CVs (>43%) both the scaling cap and the GMR-restriction help to maintain the TIE <0.05).





Outlook.

Utopia

— Agencies collect CV_{wR} from submitted studies. Pool them, adjust for designs / degrees of freedom. The EMA publishes a fixed acceptance range in the product-specific guidance. No need for replicate studies any more. 2×2×2 crossovers evaluated by ABE would be sufficient.

Halfbaked

- Hope [sic] that e.g., Bonferroni preserves the consumer risk. Still apply ABEL, but with a 95% CI (α 0.025).
- Drawback: Loss of power, substantial increase in sample sizes.

Proposal

— Iteratively adjust α based on the study's CV_{wR} and sample size — in such a way that the consumer risk is preserved (Labes/Schütz 2016).

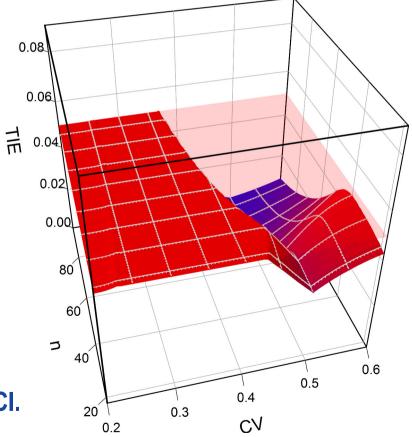




Previous example

- Algorithm
 - Assess the TIE for the nominal α 0.05.
 - If the TIE \leq 0.05, stop.
 - Otherwise adjust α (downwards) until the TIE = 0.05.
 - At CV_{wR} 30% (dependent on the sample size) α_{adj} is 0.0273 0.0300;

 \rightarrow use a 94.00 – 94.54% CI.

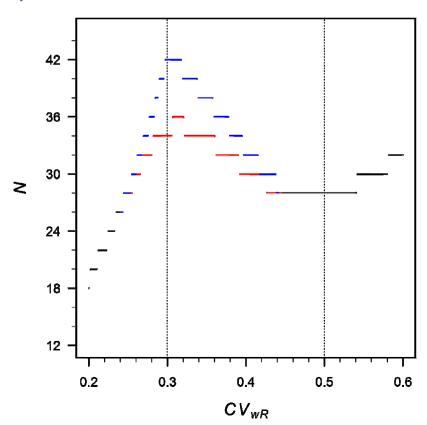






Potential impact on the sample size.

- Example: RTRT | TRTR, θ_0 0.90, target power 0.80.
 - Moderate in the critical region (— —).
 - CV_{WR} 30%: 36 \rightarrow 42 (+17%);
 - CV_{wR} 35%: 34 \rightarrow 38 (+12%);
 - CV_{WR} 40%: 30 \rightarrow 32 (+7%).
 - None outside (—).





Example (RTRT | TRTR, expected CV_{wR} 35%, θ_0 0.90, target power 0.80); R package PowerTOST (\geq 1.3-3).

Estimate the sample size.

Estimate the empiric TIE for this study.

```
UL <- scabel(CV=0.35)[["upper"]] # scaled limit (1.2948 for CVwR 0.35) power.scabel(CV=0.35, theta0=UL, n=34, design="2x2x4", nsims=1e6) [1] 0.065566
```

• Iteratively adjust α .

```
scabel.ad(CV=0.35, n=34, design="2x2x4")
++++++++ scaled (widened) ABEL ++++++++
        iteratively adjusted alpha
CVWR 0.35, n(i) 17|17 (N 34)
Nominal alpha
                            : 0.05
                : 0.9000
Null (true) ratio
Regulatory settings : EMA (ABEL)
Empiric TIE for alpha 0.0500 : 0.06557
                     : 0.812
Power for theta0 0.900
Iteratively adjusted alpha : 0.03630
Empiric TIE for adjusted alpha: 0.05000
Power for theta0 0.900
                           : 0.773
```





 Optionally compensate for the loss in power (0.812 → 0.773) by increasing the sample size:

```
sampleN.scABEL.ad(CV=0.35, theta0=0.90, targetpower=0.80, design="2x2x4")
  ++++++++ scaled (widened) ABEL ++++++++
              Sample size estimation
          for iteratively adjusted alpha
  Study design: 2x2x4 (RTRT|TRTR)
  Expected CVwR 0.35
  Nominal alpha : 0.05
  Null (true) ratio : 0.9000
  Target power : 0.8
  Regulatory settings: EMA (ABEL)
  Switching CVwR : 30%
  Regulatory constant: 0.760
  Expanded limits : 0.7723...1.2948
  Upper scaling cap : CVwR 0.5
  PE constraints : 0.8000...1.2500
  n 38, adj. alpha: 0.03610 (power 0.8100), TIE: 0.05000
- n 34 \rightarrow 38 (+12%), power 0.773 \rightarrow 0.810, lpha_{adj} 0.0363 \rightarrow 0.0361.
```





8.0

0.6

0.4

0.2

GMR, target power

0.7 0.75 0.8 0.85 0.9 0.95 1

expected power

Side Effect

Allowing ABEL only for C_{max} .

- Some drugs show high variability in AUC as well.
 - Since in such a case the sample size will be mandated by AUC, products with high deviations in C_{max} will be approved.
 - Example: CV_{wR} 90% (C_{max}), 60% (AUC), θ_0 0.90, target power 80% \rightarrow the study is 'overpowered' for C_{max} ; C_{max} -GMRs of [0.846–1.183] will pass BE. Really desirable?
 - With the FDA's RSABE the study could be performed in only 34 subjects...

AUC (CV = 0.6)
ABE

1
1
1
0.8
0.8

0.6

0.4

0.6

0.4

0.2

8.0

0.6

0.4

0.2

GMR, target power

0.9

BE limits: 0.8000...1.2500

expected power

0.85

ABEL (EMA): design RTRT|TRTR, target power = 0.8, n = 138 (sample size dependent on AUC)

GMR GMR

Inflation of the Type I Error in Referencescaled Average Bioequivalence



Thank You! Open Questions?



Helmut Schütz BEBAC

Consultancy Services for Bioequivalence and Bioavailability Studies 1070 Vienna, Austria

helmut.schuetz@bebac.at





To bear in Remembrance...

The fundamental cause of trouble in the world today is that the stupid are cocksure while the intelligent are full of doubt.

Bertrand Russell





100% of all disasters are failures of design, not analysis.

Ronald G. Marks

My definition of an expert in any field is a person who knows enough about what's really going on to be scared.



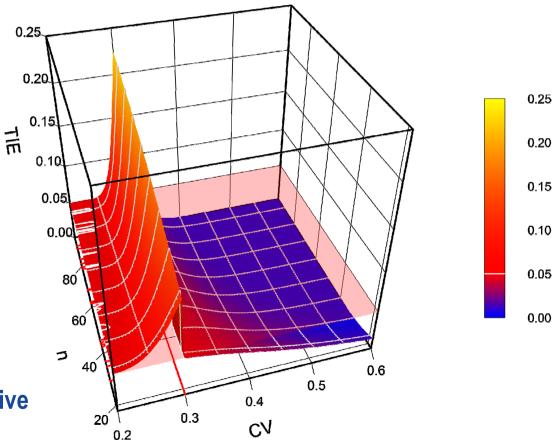
Phillip J. Plauger



Backup

Example for the FDA's RSABE

- RTRT | TRTR
 sample size 18 96
 CV_{wR} 20% 60%
 - TIE_{max} 0.2245.
 - Relative increase of the consumer risk 349%!
 - TIE more dependent on the sample size than in ABEL.
 - However, no inflation of the TIE for $CV_{wR} > 30\%$; RSABE is very conservative for 'true' HVD(P)s.



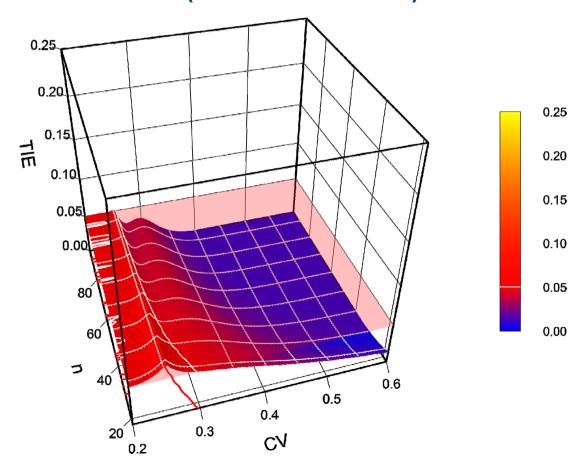




Backup

FDA's desired consumer risk model (Davit et al. 2012)

- Previous example
 - TIE assessed not at the scaled limits but
 - at 1.25 if CV_{wR} ≤25.4%
 - at $e^{0.893 \cdot \sigma_{WR}}$ otherwise.
 - TIE_{max} 0.0668.
 - Lászlo Endrényi: "Hocus pocus!"







References

- Schuirmann DJ. A Comparison of the Two One-Sided Tests Procedure and the Power Approach for Assessing the Equivalence of Average Bioavailability.

 J Pharmacokinet Biopharm. 1987; 15(6): 657–80.
- Tóthfalusi L et al. Evaluation of the Bioequivalence of Highly-Variable Drugs and Drug Products. Pharm Res. 2001;18(6): 728–33.
- Chow S-C, Shao J, Wang H. *Individual bioequivalence testing under 2×3 designs*. Stat Med. 2002; 21(5): 629–48. DOI 10.1002/sim.1056
- Tóthfalusi L, Endrényi L. Limits for the Scaled Average Bioequiva-lence of Highly Variable Drugs and Drug Products. Pharm Res. 2003; 20(3): 382–9.
- Willavize SA, Morgenthien EA. Comparison of models for average bioequivalence in replicated crossover designs. Pharm Stat. 2006; 5(3): 201–11. DOI 10.1002/pst.212
- Wolfsegger MJ, Jaki T. Simultaneous confidence intervals by iteratively adjusted alpha for relative effects in the one-way layout. Stat Comput. 2006; 16(1): 15–23. DOI 10.1007/s11222-006-5197-1
- Endrényi L, Tóthfalusi L. Regulatory Conditions for the Determination of Bioequivalence of Highly Variable Drugs. J Pharm Pharmaceut Sci. 2009; 12(1): 138–49.
- European Medicines Agency, CHMP. Guideline on the Investigation of Bioequivalence. London: 2010 Jan 20.
- http://www.ema.europa.eu/docs/en_GB/document_library/Scientific_guideline/2010/01/WC500070039.pdf
- Tóthfalusi L, Endrényi L. Sample Sizes for Designing Bioequivalence Studies for Highly Variable Drugs. J Pharm Pharmaceut Sci. 2011; 15(1): 73–84.
- Davit BM et al. Implementation of a Reference-Scaled Average Bioequivalence Approach for Highly Variable Generic Drug Products by the US Food and Drug Administration. AAPS J. 2012; 14(4): 915–24. DOI 10.1208/s12248-012-9406-x
- Patterson SD, Jones B. *Viewpoint: observations on scaled average bioequivalence.* Pharmaceut Stat. 2012; 11(1): 1–7. <u>DOI 10.1002/pst.498</u>
- Labes D. RSABE/ABEL: 'alpha' of scaled ABE? In: Bioequivalence and Bioavailability Forum [Internet]. Vienna: BEBAC; 2013 Mar 15. http://forum.bebac.at/mix_entry.php?id=10202

- European Medicines Agency, CHMP. Questions & Answers: positions on specific questions addressed to the Pharmacokinetics Working Party (PKWP). London; 2015 Nov 19.
- http://www.ema.europa.eu/docs/en_GB/document_library/Scientific_guideline/2009/09/WC500002963.pdf
- Wonnemann M, Frömke C, Koch A. *Inflation of the Type I Error: Investigations on Regulatory Recommendations for Bioequivalence of Highly Variable Drugs*. Pharm Res. 2015; 32(1): 135–43. DOI 10.1007/s11095-014-1450-z
- Labes D, Schütz H, Lang B. PowerTOST: Power and Sample size based on Two One-Sided t-Tests (TOST) for (Bio)Equivalence Studies. R package version 1.4-2. 2016. https://cran.r-project.org/package=PowerTOST
- Muñoz J, Daniel Alcaide D, Ocaña J. Consumer's risk in the EMA and FDA regulatory approaches for bioequivalence in highly variable drugs. Stat Med. 2016; 35(12): 1933–43. DOI 10.1002/sim.6834
- Labes D, Schütz H. *Inflation of Type I Error in the Evaluation of Scaled Average Bioequivalence, and a Method for its Control.* Pharm Res. Epub ahead of print: 1 August 2016. <u>DOI 10.1007/s11095-016-2006-1</u>

