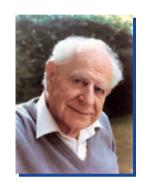




#### To bear in Remembrance...

Whenever a theory appears to you as the only possible one, take this as a sign that you have neither understood the theory nor the problem which it was intended to solve.



Karl R. Popper

Even though it's *applied* science we're dealin' with, it still is – *science*!



Leslie Z. Benet



### Bioequivalence

#### **BE** = (Desired) result of a comparative bioavailability study.

- Generally only for extravascular routes. Exceptions for IV:
  - Excipients which may interact with the API (complex formation).
  - Case-by-case: Liposomal formulations, emulsions.
- Same active substance.
  - Focus on the 'core' API
     (different salts, esters, isomers, complexes contain the same API).
- Same molar dose.
- Clinically not relevant difference: △ 20% (NTIDs 10%, HVD(P)s >20%).
- 100(1 2 $\alpha$ ) confidence interval of PK-metrics within [1  $\Delta$ , (1  $\Delta$ )<sup>-1</sup>].
  - $AUC_{0-t}$  (extent of BA)
  - $C_{max}$  (rate of BA)
  - $-t_{max}$ ,  $AUC_{0-\tau}$ ,  $C_{max,ss}$ ,  $C_{min,ss}$ ,  $C_{\tau,ss}$ , %PTF, partial AUCs, ...



## **Study Designs**

#### ≥1 Test Treatment(s) compared to ≥1 Reference Treatment(s).

- Parallel Group(s)
  - APIs with (very) long half-lives.
  - Studies in patients.
- Crossover
  - Preferred design in BE.
  - More powerful than parallel (based on within subject variance).
- Replicate crossover
  - At least one treatment is administered more than once.
  - Allows estimation of within subject variance of treatment(s).
  - Required for reference-scaling.



### **Study Designs**

The more 'sophisticated' a design is, the more information can be extracted.

Hierarchy of designs:

```
Full replicate (RTRT | TRTR or RTR | TRT) →
Partial replicate (RRT | RTR | TRR) →
2×2×2 crossover (RT | TR) →
Parallel (R | T)
```

Variances which can be estimated:

Parallel: total variance (between + within subjects)

2×2×2 crossover: + between, within subjects 🖈

Partial replicate: + within subjects (of R) *→* 

Full replicate: + within subjects (of R and T) *→* 



## **Assumptions**

#### All models rely on assumptions.

- Bioequivalence as a surrogate for therapeutic equivalance.
  - Studies in healthy volunteers in order to minimize variability (i.e., lower sample sizes than in patients).
  - Current emphasis on in vivo release ('human dissolution apparatus').
- Concentrations in the sample matrix reflect concentrations at the target receptor site.
  - In the strict sense only valid in steady state.
  - In vivo similarity in healthy volunteers can be extrapolated to the patient population(s).
- $f = \mu_T / \mu_R$  assumes that
  - $-D_T = D_R$  and
  - inter-occasion clearances are constant.



### **Assumptions**

#### All models rely on assumptions.

- Log-transformation allows for additive effects required in ANOVA.
- No carry-over effect in the model of crossover studies.
  - Cannot be statistically adjusted.
  - Has to be avoided by design (suitable washout).
  - Shown to be a statistical artifact in meta-studies.
  - Exception: Endogenous compounds (biosimilars!)
- Between- and within-subject errors are independently and normally distributed about unity with variances  $\sigma_s$  and  $\sigma_e$ .
  - If the reference formulation shows higher variability than the test,
     the 'good' test will be penalized for the 'bad' reference.
- All observations made on different subjects are independent.
  - No monocygotic twins or triplets in the study!

#### **Excursion 1**

#### Type I Error.

- In BE the Null Hypothesis is inequivalence.
  - TIE = Probability of falsely rejecting the Null (i.e., claiming BE).
  - Can be calculated for the nominal significance level ( $\alpha$ ) assuming a PE at one of the limits of the acceptance range.
    - Example: 2×2×2 crossover, CV 20%, n 20,  $\alpha$  0.05,  $\theta_0$  1.25.

```
library(PowerTOST)
AL <- c(0.80, 1.25) # common range for ABE
power.TOST(CV=0.20, n=20, alpha=0.05, theta0=AL[1])
[1] 0.0499999
power.TOST(CV=0.20, n=20, alpha=0.05, theta0=AL[2])
[1] 0.0499999
```

TOST is not a uniformly most powerful test.

```
power.TOST(CV=0.20, n=12, alpha=0.05, theta0=AL[2]) [1] 0.04976374
```

However, the TIE never exceeds its nominal level.

```
power.TOST(CV=0.20, n=120, alpha=0.05, theta0=AL[2]) [1] 0.05
```



#### **Excursion 1**

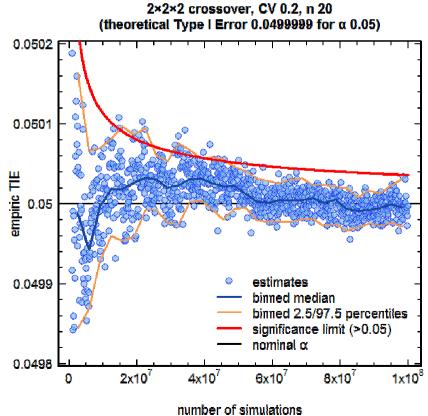
#### Type I Error.

Alternatively perform simulations to obtain an empiric TIE.

```
power.TOST.sim(CV=0.20, n=20, alpha=0.05, theta0=AL[2], nsims=1e8)
[1] 0.0499970

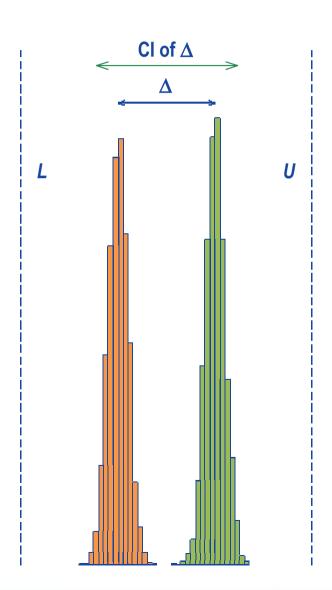
2*2*2 crossover, C
```

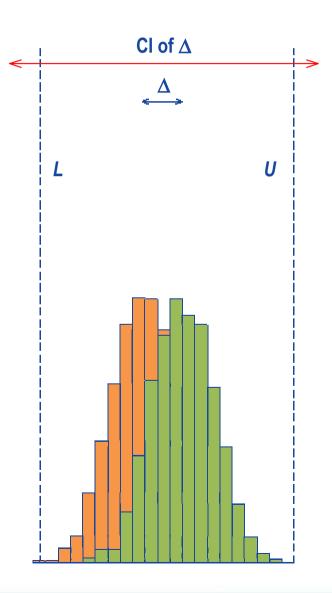
 In other settings (e.g., Two-Stage Designs or reference-scaled ABE) analytical solutions for power (and therefore, the TIE) are not possible.





## Highly Variable Drugs / Drug Products





## Counterintuitive concept of BE:

Two formulations with a large difference in means are declared bioequivalent if variances are low, but not BE – even if the difference is quite small – due to high variability.

Modified from Tothfálusi et al. (2009), Fig. 1



## It may be almost impossible to demonstrate BE with a reasonable sample size.

- Reference-scaling (*i.e.*, widening the acceptance range based of the variability of the reference) in 2010 introduced by the FDA and EMA and in 2016 by Health Canada.
  - Requires a replicate design, where at least the reference product is administered twice.
  - Smaller sample sizes compared to a standard 2×2×2 design but outweighed by increased number of periods.
  - Similar total number of individual treatments.
  - Any replicate design can be evaluated for 'classical' (unscaled) Average Bioequivalence (ABE) as well. Switching  $CV_{wR}$  30%:
    - FDA: AUC and  $C_{max}$
    - EMA:  $C_{max}$ ; MR products additionally:  $C_{ss,min}$ ,  $C_{ss,r}$ , partial AUCs
    - Health Canada: AUC



#### Models (in log-scale).

- ABE Model:
  - A difference  $\triangle$  of ≤20% is considered to be clinically not relevant.
  - The limits [L, U] of the acceptance range are fixed to  $log(1 \Delta) = log((1 \Delta)^{-1})$  or  $L \sim -0.2231$  and  $U \sim +0.2231$ .
  - The consumer risk is fixed with 0.05. BE is concluded if the  $100(1 2\alpha)$  confidence interval lies entirely within the acceptance range.

$$-\theta_{A} \leq \mu_{T} - \mu_{R} \leq +\theta_{A}$$

- SABEL Model:
  - Switching condition  $\theta_S$  is derived from the regulatory standardized variation  $\sigma_0$  (proportionality between acceptance limits in log-scale and  $\sigma_{wR}$  in the highly variable region).

$$-\theta_{S} \leq \frac{\mu_{T} - \mu_{R}}{\sigma_{WR}} \leq +\theta_{S}$$

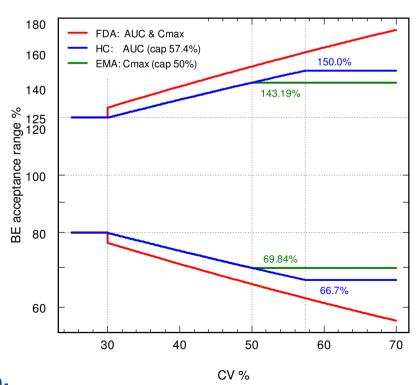


#### Regulatory Approaches.

• Bioequivalence limits derived from  $\sigma_{\!_{0}}$  and  $\sigma_{\!_{wR}}$ 

$$\theta_{S} = \frac{\log(1.25)}{\sigma_{0}}, [L,U] = e^{\pm\theta_{S}\cdot\sigma_{WR}}$$

- FDA
  - Scaling  $\sigma_{wR}$  0.25 ( $\theta_{S}$  0.893) but applicable at  $CV_{wR} \ge 30\%$ .
  - Discontinuity at  $CV_{wR}$  30%.
- EMA
  - Scaling  $\sigma_0$  0.2936 ( $\theta_S$  0.760).
  - Upper cap at  $CV_{wR}$  50%.
- Health Canada
  - Like EMA but upper cap at  $CV_{wR}$  57.4%.





#### The EMA's Approach.

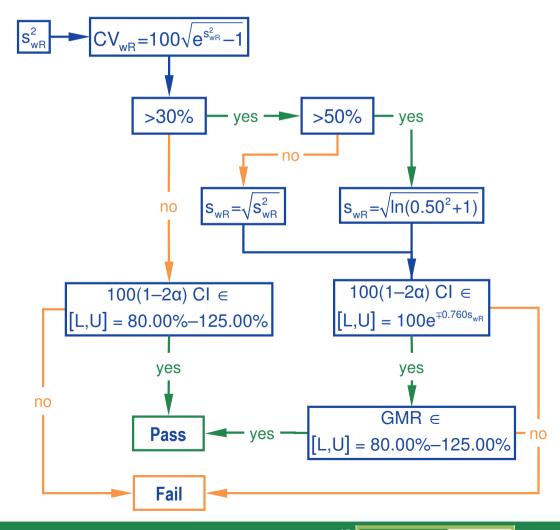
- Average Bioequivalence with Expanding Limits ABEL (crippled from Endrényi and Tóthfalusi 2009).
  - Justification that the widened acceptance range is clinically not relevant (important – different to the FDA).
  - Assumes identical variances of T and R [sic] like in a 2×2×2.
  - All fixed effects model according to the Q&A-document preferred.
  - Mixed-effects model (allowing for unequival variances) is 'not compatible with CHMP guideline'...
  - Scaling limited at a maximum of  $CV_{wR}$  50% (i.e., to 69.84 143.19%).
  - GMR within 0.8000 1.2500.
  - Demonstration that  $CV_{wR} > 30\%$  is not caused by outliers (box plots of studentized intra-subject residuals?)...
  - ≥12 subjects in sequence RTR of the 3-period full replicate design.



#### The EMA's Approach.

- Decision Scheme.
  - The Null Hypothesis is specified in the face of the data.
  - Acceptance limits themselves become random variables.
  - Type I Error (consumer risk) might be inflated.







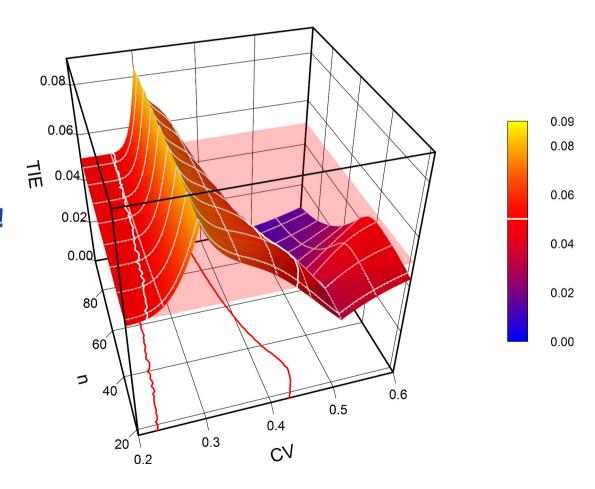
#### Assessing the Type I Error (TIE).

- TIE = falsely concluding BE at the limits of the acceptance range. In ABE the TIE is  $\leq 0.05$  at 0.80 and  $\leq 0.05$  at 1.25.
- Due to the decision scheme no direct calculation of the TIE at the scaled limits is possible;
  - → extensive simulations required (10<sup>6</sup> BE studies mandatory).
- Inflation of the TIE suspected. (Chow et al. 2002, Willavazie & Morgenthien 2006, Chow & Liu 2009, Patterson & Jones 2012).
- Confirmed.
  - EMA's ABEL
     (Tóthfalusi & Endrényi 2009, BEBA-Forum 2013, Wonnemann et al. 2015, Muñoz et al. 2016, Labes & Schütz 2016).
  - FDA's RSABE
     (Tóthfalusi & Endrényi 2009, BEBA-Forum 2013, Muñoz et al. 2016).



#### **Example for ABEL**

- RTRT | TRTR
   sample size 18 96
   CV<sub>wR</sub> 20% 60%
  - TIE<sub>max</sub> 0.0837.
  - Relative increase of the consumer risk 67%!





#### What is going on here?

SABE is stated in model parameters ...

$$-\theta_{S} \leq \frac{\mu_{T} - \mu_{R}}{\sigma_{WR}} \leq +\theta_{S}$$

- ... which are unknown.
- Only their estimates (GMR,  $s_{wR}$ ) are accessible in the actual study.
- At  $CV_{wR}$  30% the decision to scale will be wrong in ~50% of cases.
- If moving away from 30% the chances of a wrong decision decrease and hence, the TIE.
- At high CVs (>43%) both the scaling cap and the GMR-restriction help to maintain the TIE <0.05).</li>



#### Outlook.

#### Utopia

— Agencies collect  $CV_{wR}$  from submitted studies. Pool them, adjust for designs / degrees of freedom. The EMA publishs a fixed acceptance range in the product-specific guidance. No need for replicate studies any more. 2×2×2 crossovers evaluated by ABE would be sufficient.

#### Halfbaked

- Hope [sic] that e.g., Bonferroni preserves the consumer risk. Still apply ABEL, but with a 95% CI ( $\alpha$ 0.025).
- Drawback: Loss of power, substantial increase in sample sizes.

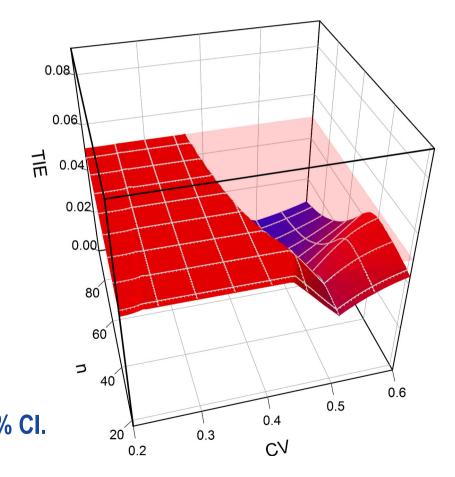
#### Proposal

— Iteratively adjust  $\alpha$  based on the study's  $CV_{wR}$  and sample size – in such a way that the consumer risk is preserved.



#### **Previous example**

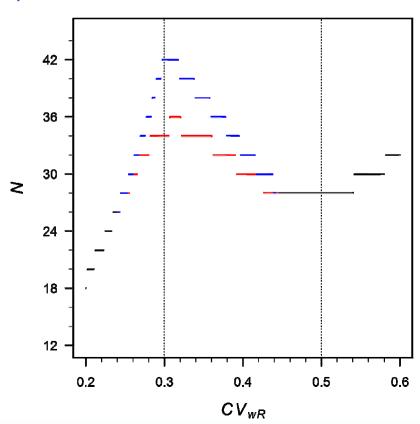
- Algorithm
  - Assess the TIE for the nominal  $\alpha$  0.05.
  - If the TIE  $\leq$  0.05, stop.
  - Otherwise adjust  $\alpha$  (downwards) until the TIE = 0.05.
  - At  $CV_{wR}$  30% (dependent on the sample size)  $\alpha_{adj}$  is 0.0273 - 0.0300; -> use a 94.00 - 94.54% CI.





#### Potential impact on the sample size.

- Example: RTRT | TRTR,  $\theta_0$  0.90, target power 0.80.
  - Moderate in the critical region (— —).
    - $CV_{WR}$  30%: 36  $\rightarrow$  42 (+17%);
    - $CV_{WR}$  35%: 34  $\rightarrow$  38 (+12%);
    - $CV_{WR}$  40%: 30  $\rightarrow$  32 ( +7%).
  - None outside (—).





## Example (RTRT | TRTR, expected $CV_{wR}$ 35%, $\theta_0$ 0.90, target power 0.80); R package PowerTOST ( $\geq$ 1.3-3).

Estimate the sample size.

Estimate the empiric TIE for this study.

```
UL <- scabel(CV=0.35)[["upper"]] # scaled limit (1.2948 for CVwR 0.35) power.scabel(CV=0.35, theta0=UL, n=34, design="2x2x4", nsims=1e6) [1] 0.065566
```

• Iteratively adjust  $\alpha$ .

```
scabel.ad(CV=0.35, n=34, design="2x2x4")
++++++++ scaled (widened) ABEL ++++++++
        iteratively adjusted alpha
CVwR 0.35, n(i) 17|17 (N 34)
Nominal alpha
                             : 0.05
Null (true) ratio
                            : 0.9000
Regulatory settings : EMA (ABEL)
Empiric TIE for alpha 0.0500 : 0.06557
Power for theta0 0.900
                           : 0.812
Iteratively adjusted alpha : 0.03630
Empiric TIE for adjusted alpha: 0.05000
Power for theta0 0.900
                            : 0.773
```



 Optionally compensate for the loss in power (0.812 → 0.773) by increasing the sample size:

```
sampleN.scABEL.ad(CV=0.35, theta0=0.90, targetpower=0.80, design="2x2x4")
  ++++++++ scaled (widened) ABEL ++++++++
              Sample size estimation
          for iteratively adjusted alpha
  Study design: 2x2x4 (RTRT|TRTR)
  Expected CVwR 0.35
  Nominal alpha
                     : 0.05
  Null (true) ratio : 0.9000
  Target power : 0.8
  Regulatory settings: EMA (ABEL)
  Switching CVwR : 30%
  Regulatory constant: 0.760
  Expanded limits : 0.7723...1.2948
  Upper scaling cap : CVwR 0.5
  PE constraints : 0.8000...1.2500
  n 38, adj. alpha: 0.03610 (power 0.8100), TIE: 0.05000
- n 34 \rightarrow 38 (+12%), power 0.773 \rightarrow 0.810, lpha_{adi} 0.0363 \rightarrow 0.0361.
```

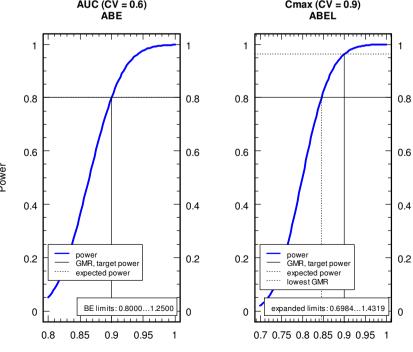


#### **Excursion 2**

### 'Side effect' of allowing ABEL only for $C_{max}$ .

- Some drugs show high variability in AUC as well.
  - Since in such a case the sample size will be mandated by AUC, products with high deviations in  $C_{max}$  will be approved.
  - Example:  $CV_{wR}$  90% ( $C_{max}$ ), 60% (AUC),  $\theta_0$  0.90, target power 80%  $\rightarrow$  the study is 'overpowered' for  $C_{max}$ ;  $C_{max}$ -GMRs of [0.846–1.183] will pass BE. Really desirable?
  - With the FDA's RSABE the study could be performed in only 34 subjects...





GMR GMR

## Inflation of the Type I Error in Referencescaled Average Bioequivalence



# Thank You! Open Questions?



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#### To bear in Remembrance...

The fundamental cause of trouble in the world today is that the stupid are cocksure while the intelligent are full of doubt.

Bertrand Russell





100% of all disasters are failures of design, not analysis.

Ronald G. Marks

My definition of an expert in any field is a person who knows enough about what's really going on to be scared.



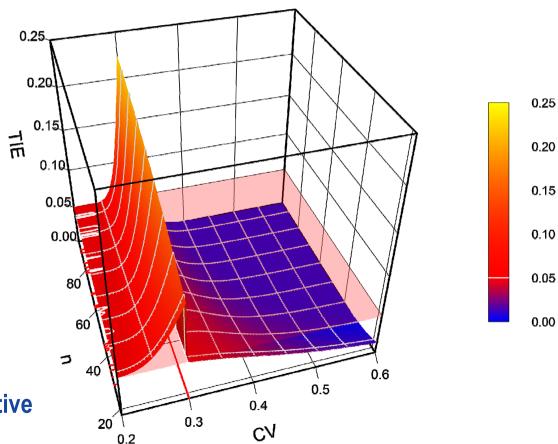
Phillip J. Plauger





#### **Example for the FDA's RSABE**

- RTRT | TRTR
   sample size 18 96
   CV<sub>wR</sub> 20% 60%
  - TIE<sub>max</sub> 0.2245.
  - Relative increase of the consumer risk 349%!
  - TIE more dependent on the sample size than in ABEL.
  - However, no inflation of the TIE for CV<sub>wR</sub> >30%; RSABE is very conservative for 'true' HVD(P)s.

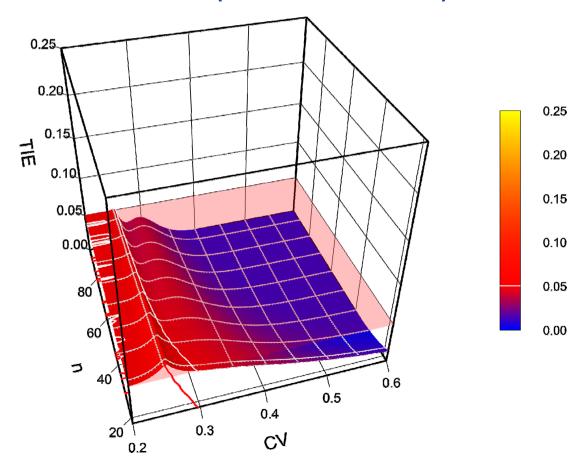




## Backup

#### "FDA's desired consumer risk model" (Davit et al. 2012)

- Previous example
  - TIE assessed not at the scaled limits but
    - at 1.25 if CV<sub>wR</sub> ≤25.4%
    - at  $e^{0.893 \cdot \sigma_{WR}}$  otherwise.
  - TIE<sub>max</sub> 0.0668.
  - Lászlo Endrényi: "Hocus pocus!"



## BE ·

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